Exercises for the tutorials: 1, 3.
The other exercises are for self-study and exam preparation. All material is examinable unless otherwise mentioned.

## Exercise 1. Predictive distributions for hidden Markov models

For the hidden Markov model

$$
p\left(h_{1: d}, v_{1: d}\right)=p\left(v_{1} \mid h_{1}\right) p\left(h_{1}\right) \prod_{i=2}^{d} p\left(v_{i} \mid h_{i}\right) p\left(h_{i} \mid h_{i-1}\right)
$$

assume you have observations for $v_{i}, i=1, \ldots, u<d$.
(a) Use message passing to compute $p\left(h_{t} \mid v_{1: u}\right)$ for $u<t \leq d$. For the sake of concreteness, you may consider the case $d=6, u=2, t=4$.
(b) Use message passing to compute $p\left(v_{t} \mid v_{1: u}\right)$ for $u<t \leq d$. For the sake of concreteness, you may consider the case $d=6, u=2, t=4$.

## Exercise 2. Viterbi algorithm

For the hidden Markov model

$$
p\left(h_{1: t}, v_{1: t}\right)=p\left(v_{1} \mid h_{1}\right) p\left(h_{1}\right) \prod_{i=2}^{t} p\left(v_{i} \mid h_{i}\right) p\left(h_{i} \mid h_{i-1}\right)
$$

assume you have observations for $v_{i}, i=1, \ldots, t$. Use the max-sum algorithm to derive an iterative algorithm to compute

$$
\begin{equation*}
\hat{\mathbf{h}}=\underset{h_{1}, \ldots, h_{t}}{\operatorname{argmax}} p\left(h_{1: t} \mid v_{1: t}\right) \tag{1}
\end{equation*}
$$

Assume that the latent variables $h_{i}$ can take $K$ different values, e.g. $h_{i} \in\{0, \ldots, K-1\}$. The resulting algorithm is known as Viterbi algorithm.

## Exercise 3. Forward filtering backward sampling for hidden Markov models

Consider the hidden Markov model specified by the following DAG.


We assume that have already run the alpha-recursion (filtering) and can compute $p\left(h_{t} \mid v_{1: t}\right)$ for all $t$. The goal is now to generate samples $p\left(h_{1}, \ldots, h_{n} \mid v_{1: n}\right)$, i.e. entire trajectories $\left(h_{1}, \ldots, h_{n}\right)$
from the posterior. Note that this is not the same as sampling from the $n$ filtering distributions $p\left(h_{t} \mid v_{1: t}\right)$. Moreover, compared to the Viterbi algorithm, the sampling approach generates samples from the full posterior rather than just returning the most probable state and its corresponding probability.
(a) Show that $p\left(h_{1}, \ldots, h_{n} \mid v_{1: n}\right)$ forms a first-order Markov chain.
(b) Since $p\left(h_{1}, \ldots, h_{n} \mid v_{1: n}\right)$ is a first-order Markov chain, it suffices to determine $p\left(h_{t-1} \mid h_{t}, v_{1: n}\right)$, the probability mass function for $h_{t-1}$ given $h_{t}$ and all the data $v_{1: n}$. Use message passing to show that

$$
\begin{equation*}
p\left(h_{t-1}, h_{t} \mid v_{1: n}\right) \propto \alpha\left(h_{t-1}\right) \beta\left(h_{t}\right) p\left(h_{t} \mid h_{t-1}\right) p\left(v_{t} \mid h_{t}\right) \tag{2}
\end{equation*}
$$

(c) Show that $p\left(h_{t-1} \mid h_{t}, v_{1: n}\right)=\frac{\alpha\left(h_{t-1}\right)}{\alpha\left(h_{t}\right)} p\left(h_{t} \mid h_{t-1}\right) p\left(v_{t} \mid h_{t}\right)$.

We thus obtain the following algorithm to generate samples from $p\left(h_{1}, \ldots, h_{n} \mid v_{1: n}\right)$ :

1. Run the alpha-recursion (filtering) to determine all $\alpha\left(h_{t}\right)$ forward in time for $t=$ $1, \ldots, n$.
2. Sample $h_{n}$ from $p\left(h_{n} \mid v_{1: n}\right) \propto \alpha\left(h_{n}\right)$
3. Go backwards in time using

$$
\begin{equation*}
p\left(h_{t-1} \mid h_{t}, v_{1: n}\right)=\frac{\alpha\left(h_{t-1}\right)}{\alpha\left(h_{t}\right)} p\left(h_{t} \mid h_{t-1}\right) p\left(v_{t} \mid h_{t}\right) \tag{3}
\end{equation*}
$$

to generate samples $h_{t-1} \mid h_{t}, v_{1: n}$ for $t=n, \ldots, 2$.
This algorithm is known as forward filtering backward sampling (FFBS).

## Exercise 4. Prediction exercise

Consider a hidden Markov model with three visibles $v_{1}, v_{2}, v_{3}$ and three hidden variables $h_{1}, h_{2}, h_{3}$ which can be represented with the following factor graph:


This question is about computing the predictive probability $p\left(v_{3}=1 \mid v_{1}=1\right)$.
(a) The factor graph below represents $p\left(h_{1}, h_{2}, h_{3}, v_{2}, v_{3} \mid v_{1}=1\right)$. Provide an equation that defines $\phi_{A}$ in terms of the factors in the factor graph above.

(b) Assume further that all variables are binary, $h_{i} \in\{0,1\}, v_{i} \in\{0,1\}$; that $p\left(h_{1}=1\right)=0.5$, and that the transition and emission distributions are, for all $i$, given by:

| $p\left(h_{i+1} \mid h_{i}\right)$ | $h_{i+1}$ | $h_{i}$ |
| :--- | :--- | :--- |
| 0 | 0 | 0 |
| 1 | 1 | 0 |
| 1 | 0 | 1 |
| 0 | 1 | 1 |


| $p\left(v_{i} \mid h_{i}\right)$ | $v_{i}$ | $h_{i}$ |
| :--- | :--- | :--- |
| 0.6 | 0 | 0 |
| 0.4 | 1 | 0 |
| 0.4 | 0 | 1 |
| 0.6 | 1 | 1 |

Compute the numerical values of the factor $\phi_{A}$.
(d) Denote the message from variable node $h_{2}$ to factor node $p\left(h_{3} \mid h_{2}\right)$ by $\alpha\left(h_{2}\right)$. Use message passing to compute $\alpha\left(h_{2}\right)$ for $h_{2}=0$ and $h_{2}=1$. Report the values of any intermediate messages that need to be computed for the computation of $\alpha\left(h_{2}\right)$.
(e) With $\alpha\left(h_{2}\right)$ defined as above, use message passing to show that the predictive probability $p\left(v_{3}=1 \mid v_{1}=1\right)$ can be expressed in terms of $\alpha\left(h_{2}\right)$ as

$$
\begin{equation*}
p\left(v_{3}=1 \mid v_{1}=1\right)=\frac{x \alpha\left(h_{2}=1\right)+y \alpha\left(h_{2}=0\right)}{\alpha\left(h_{2}=1\right)+\alpha\left(h_{2}=0\right)} \tag{4}
\end{equation*}
$$

and report the values of $x$ and $y$.
(f) Compute the numerical value of $p\left(v_{3}=1 \mid v_{1}=1\right)$.

## Exercise 5. Hidden Markov models and change of measure

We take here a change of measure perspective on the alpha-recursion.
Consider the following directed graph for a hidden Markov model where the $y_{i}$ correspond to observed (visible) variables and the $x_{i}$ to unobserved (hidden/latent) variables.


The joint model for $\mathbf{x}=\left(x_{1}, \ldots, x_{n}\right)$ and $\mathbf{y}=\left(y_{1}, \ldots, y_{n}\right)$ thus is

$$
\begin{equation*}
p(\mathbf{x}, \mathbf{y})=p\left(x_{1}\right) \prod_{i=2}^{n} p\left(x_{i} \mid x_{i-1}\right) \prod_{i=1}^{n} p\left(y_{i} \mid x_{i}\right) \tag{5}
\end{equation*}
$$

(a) Show that

$$
\begin{equation*}
p\left(x_{1}, \ldots, x_{n}, y_{1}, \ldots, y_{t}\right)=f_{1}\left(x_{1}\right) \prod_{i=2}^{n} f_{i}\left(x_{i} \mid x_{i-1}\right) \prod_{i=1}^{t} p\left(y_{i} \mid x_{i}\right) \tag{6}
\end{equation*}
$$

for $t=0, \ldots, n$. We take the case $t=0$ to correspond to $p\left(x_{1}, \ldots, x_{n}\right)$,

$$
\begin{equation*}
p\left(x_{1}, \ldots, x_{n}\right)=f_{1}\left(x_{1}\right) \prod_{i=2}^{n} f_{i}\left(x_{i} \mid x_{i-1}\right) \tag{7}
\end{equation*}
$$

(b) Show that $p\left(x_{1}, \ldots, x_{n} \mid y_{1}, \ldots, y_{t}\right), t=0, \ldots, n$, factorises as

$$
\begin{equation*}
p\left(x_{1}, \ldots, x_{n} \mid y_{1}, \ldots, y_{t}\right) \propto p\left(x_{1}\right) \prod_{i=2}^{n} p\left(x_{i} \mid x_{i-1}\right) \prod_{i=1}^{t} g_{i}\left(x_{i}\right) \tag{8}
\end{equation*}
$$

where $g_{i}\left(x_{i}\right)=p\left(y_{i} \mid x_{i}\right)$ for a fixed value of $y_{i}$, and that its normalising constant $Z_{t}$ equals the likelihood $p\left(y_{1}, \ldots, y_{t}\right)$
(c) Denote $p\left(x_{1}, \ldots, x_{n} \mid y_{1}, \ldots, y_{t}\right)$ by $p_{t}\left(x_{1}, \ldots, x_{n}\right)$. The index $t \leq n$ thus indicates the time of the last $y$-variable we are conditioning on. Show the following recursion for $1 \leq t \leq n$ :

$$
\begin{array}{rll}
p_{t-1}\left(x_{1}, \ldots, x_{t}\right) & = \begin{cases}p\left(x_{1}\right) & \text { if } t=1 \\
p_{t-1}\left(x_{1}, \ldots, x_{t-1}\right) p\left(x_{t} \mid x_{t-1}\right) & \text { otherwise }\end{cases} & \text { (extension) } \\
p_{t}\left(x_{1}, \ldots, x_{t}\right) & =\frac{1}{Z_{t}} p_{t-1}\left(x_{1}, \ldots, x_{t}\right) g_{t}\left(x_{t}\right) & \text { (change of measure) } \\
Z_{t} & =\int p_{t-1}\left(x_{t}\right) g_{t}\left(x_{t}\right) \mathrm{d} x_{t} & \tag{11}
\end{array}
$$

By iterating from $t=1$ to $t=n$, we can thus recursively compute $p\left(x_{1}, \ldots, x_{n} \mid y_{1}, \ldots, y_{n}\right)$, including its normalising constant $Z_{n}$, which equals the likelihood $Z_{n}=p\left(y_{1}, \ldots, y_{n}\right)$
(d) Use the recursion above to derive the following form of the alpha recursion:

$$
\begin{align*}
p_{t-1}\left(x_{t-1}, x_{t}\right) & =p_{t-1}\left(x_{t-1}\right) p\left(x_{t} \mid x_{t-1}\right) & & \text { (extension) }  \tag{12}\\
p_{t-1}\left(x_{t}\right) & =\int p_{t-1}\left(x_{t-1}, x_{t}\right) \mathrm{d} x_{t-1} & & \text { (marginalisation) }  \tag{13}\\
p_{t}\left(x_{t}\right) & =\frac{1}{Z_{t}} p_{t-1}\left(x_{t}\right) g_{t}\left(x_{t}\right) & & \text { (change of measure) }  \tag{14}\\
Z_{t} & =\int p_{t-1}\left(x_{t}\right) g_{t}\left(x_{t}\right) \mathrm{d} x_{t} & & \tag{15}
\end{align*}
$$

with $p_{0}\left(x_{1}\right)=p\left(x_{1}\right)$.
The term $p_{t}\left(x_{t}\right)$ corresponds to $\alpha\left(x_{t}\right)$ from the alpha-recursion after normalisation. As in the lecture, we see that $p_{t-1}\left(x_{t}\right)$ is a predictive distribution for $x_{t}$ given observations until time $t-1$. Multiplying $p_{t-1}\left(x_{t}\right)$ with $g_{t}\left(x_{t}\right)$ gives the new $\alpha\left(x_{t}\right)$. In the lecture we called $g_{t}\left(x_{t}\right)=p\left(y_{t} \mid x_{t}\right)$ the "correction". We see here that the correction has the effect of a change of measure, changing the predictive distribution $p_{t-1}\left(x_{t}\right)$ into the filtering distribution $p_{t}\left(x_{t}\right)$.

## Exercise 6. Kalman filtering

We here consider filtering for hidden Markov models with Gaussian transition and emission distributions. For simplicity, we assume one-dimensional hidden variables and observables. We denote the probability density function of a Gaussian random variable $x$ with mean $\mu$ and variance $\sigma^{2}$ by $\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)$,

$$
\begin{equation*}
\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} \exp \left[-\frac{(x-\mu)^{2}}{2 \sigma^{2}}\right] \tag{16}
\end{equation*}
$$

The transition and emission distributions are assumed to be

$$
\begin{align*}
p\left(h_{s} \mid h_{s-1}\right) & =\mathcal{N}\left(h_{s} \mid A_{s} h_{s-1}, B_{s}^{2}\right)  \tag{17}\\
p\left(v_{s} \mid h_{s}\right) & =\mathcal{N}\left(v_{s} \mid C_{s} h_{s}, D_{s}^{2}\right) \tag{18}
\end{align*}
$$

The distribution $p\left(h_{1}\right)$ is assumed Gaussian with known parameters. The $A_{s}, B_{s}, C_{s}, D_{s}$ are also assumed known.
(a) Show that $h_{s}$ and $v_{s}$ as defined in the following update and observation equations

$$
\begin{align*}
h_{s} & =A_{s} h_{s-1}+B_{s} \xi_{s}  \tag{19}\\
v_{s} & =C_{s} h_{s}+D_{s} \eta_{s} \tag{20}
\end{align*}
$$

follow the conditional distributions in (17) and (18). The random variables $\xi_{s}$ and $\eta_{s}$ are independent from the other variables in the model and follow a standard normal Gaussian distribution, e.g. $\xi_{s} \sim \mathcal{N}\left(\xi_{s} \mid 0,1\right)$.
Hint: For two constants $c_{1}$ and $c_{2}, y=c_{1}+c_{2} x$ is Gaussian if $x$ is Gaussian. In other words, an affine transformation of a Gaussian is Gaussian.
The equations mean that $h_{s}$ is obtained by scaling $h_{s-1}$ and by adding noise with variance $B_{s}^{2}$. The observed value $v_{s}$ is obtained by scaling the hidden $h_{s}$ and by corrupting it with Gaussian observation noise of variance $D_{s}^{2}$.
(b) Show that

$$
\begin{equation*}
\int \mathcal{N}\left(x \mid \mu, \sigma^{2}\right) \mathcal{N}\left(y \mid A x, B^{2}\right) \mathrm{d} x \propto \mathcal{N}\left(y \mid A \mu, A^{2} \sigma^{2}+B^{2}\right) \tag{21}
\end{equation*}
$$

Hint: While this result can be obtained by integration, an approach that avoids this is as follows: First note that $\mathcal{N}\left(x \mid \mu, \sigma^{2}\right) \mathcal{N}\left(y \mid A x, B^{2}\right)$ is proportional to the joint pdf of $x$ and $y$. We can thus consider the integral to correspond to the computation of the marginal of $y$ from the joint. Using the equivalence of Equations (17)-(18) and (19)-(20), and the fact that the weighted sum of two Gaussian random variables is a Gaussian random variable then allows one to obtain the result.
(c) Show that

$$
\begin{equation*}
\mathcal{N}\left(x \mid m_{1}, \sigma_{1}^{2}\right) \mathcal{N}\left(x \mid m_{2}, \sigma_{2}^{2}\right) \propto \mathcal{N}\left(x \mid m_{3}, \sigma_{3}^{2}\right) \tag{22}
\end{equation*}
$$

where

$$
\begin{align*}
& \sigma_{3}^{2}=\left(\frac{1}{\sigma_{1}^{2}}+\frac{1}{\sigma_{2}^{2}}\right)^{-1}=\frac{\sigma_{1}^{2} \sigma_{2}^{2}}{\sigma_{1}^{2}+\sigma_{2}^{2}}  \tag{23}\\
& m_{3}=\sigma_{3}^{2}\left(\frac{m_{1}}{\sigma_{1}^{2}}+\frac{m_{2}}{\sigma_{2}^{2}}\right)=m_{1}+\frac{\sigma_{1}^{2}}{\sigma_{1}^{2}+\sigma_{2}^{2}}\left(m_{2}-m_{1}\right) \tag{24}
\end{align*}
$$

Hint: Work in the negative log domain.
(d) In the lecture, we have seen that $p\left(h_{t} \mid v_{1: t}\right) \propto \alpha\left(h_{t}\right)$ where $\alpha\left(h_{t}\right)$ can be computed recursively via the "alpha-recursion"

$$
\begin{equation*}
\alpha\left(h_{1}\right)=p\left(h_{1}\right) \cdot p\left(v_{1} \mid h_{1}\right) \quad \alpha\left(h_{s}\right)=p\left(v_{s} \mid h_{s}\right) \sum_{h_{s-1}} p\left(h_{s} \mid h_{s-1}\right) \alpha\left(h_{s-1}\right) . \tag{25}
\end{equation*}
$$

For continuous random variables, the sum above becomes an integral so that

$$
\begin{equation*}
\alpha\left(h_{s}\right)=p\left(v_{s} \mid h_{s}\right) \int p\left(h_{s} \mid h_{s-1}\right) \alpha\left(h_{s-1}\right) \mathrm{d} h_{s-1} . \tag{26}
\end{equation*}
$$

For reference, let us denote the integral by $I\left(h_{s}\right)$,

$$
\begin{equation*}
I\left(h_{s}\right)=\int p\left(h_{s} \mid h_{s-1}\right) \alpha\left(h_{s-1}\right) \mathrm{d} h_{s-1} \tag{27}
\end{equation*}
$$

In the lecture, it was pointed out that $I\left(h_{s}\right)$ is proportional to the predictive distribution $p\left(h_{s} \mid v_{1: s-1}\right)$.
For a Gaussian prior distribution for $h_{1}$ and Gaussian emission probability $p\left(v_{1} \mid h_{1}\right)$, $\alpha\left(h_{1}\right)=p\left(h_{1}\right) \cdot p\left(v_{1} \mid h_{1}\right) \propto p\left(h_{1} \mid v_{1}\right)$ is proportional to a Gaussian. We denote its mean by $\mu_{1}$ and its variance by $\sigma_{1}^{2}$ so that

$$
\begin{equation*}
\alpha\left(h_{1}\right) \propto \mathcal{N}\left(h_{1} \mid \mu_{1}, \sigma_{1}^{2}\right) \tag{28}
\end{equation*}
$$

Assuming $\alpha\left(h_{s-1}\right) \propto \mathcal{N}\left(h_{s-1} \mid \mu_{s-1}, \sigma_{s-1}^{2}\right)$ (which holds for $s=2$ ), use Equation (21) to show that

$$
\begin{equation*}
I\left(h_{s}\right) \propto \mathcal{N}\left(h_{s} \mid A_{s} \mu_{s-1}, P_{s}\right) \tag{29}
\end{equation*}
$$

where

$$
\begin{equation*}
P_{s}=A_{s}^{2} \sigma_{s-1}^{2}+B_{s}^{2} \tag{30}
\end{equation*}
$$

(e) Use Equation (22) to show that

$$
\begin{equation*}
\alpha\left(h_{s}\right) \propto \mathcal{N}\left(h_{s} \mid \mu_{s}, \sigma_{s}^{2}\right) \tag{31}
\end{equation*}
$$

where

$$
\begin{align*}
\mu_{s} & =A_{s} \mu_{s-1}+\frac{P_{s} C_{s}}{C_{s}^{2} P_{s}+D_{s}^{2}}\left(v_{s}-C_{s} A_{s} \mu_{s-1}\right)  \tag{32}\\
\sigma_{s}^{2} & =\frac{P_{s} D_{s}^{2}}{P_{s} C_{s}^{2}+D_{s}^{2}} \tag{33}
\end{align*}
$$

(f) Show that $\alpha\left(h_{s}\right)$ can be re-written as

$$
\begin{equation*}
\alpha\left(h_{s}\right) \propto \mathcal{N}\left(h_{s} \mid \mu_{s}, \sigma_{s}^{2}\right) \tag{34}
\end{equation*}
$$

where

$$
\begin{align*}
\mu_{s} & =A_{s} \mu_{s-1}+K_{s}\left(v_{s}-C_{s} A_{s} \mu_{s-1}\right)  \tag{35}\\
\sigma_{s}^{2} & =\left(1-K_{s} C_{s}\right) P_{s}  \tag{36}\\
K_{s} & =\frac{P_{s} C_{s}}{C_{s}^{2} P_{s}+D_{s}^{2}} \tag{37}
\end{align*}
$$

These are the Kalman filter equations and $K_{s}$ is called the Kalman filter gain.
(g) Explain Equation (35) in non-technical terms. What happens if the variance $D_{s}^{2}$ of the observation noise goes to zero?

