Learning for Hidden Markov Models

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Recap

- We can decompose the log marginal of any joint distribution into a sum of two terms:
 - the free energy and
 - the KL divergence between the variational and the conditional distribution
- Variational principle: Maximising the free energy with respect to the variational distribution allows us to (approximately) compute the (log) marginal and the conditional from the joint.
- We applied the variational principle to inference and learning problems.
- For parameter estimation in presence of unobserved variables: Coordinate ascent on the free energy leads to the (variational) EM algorithm.

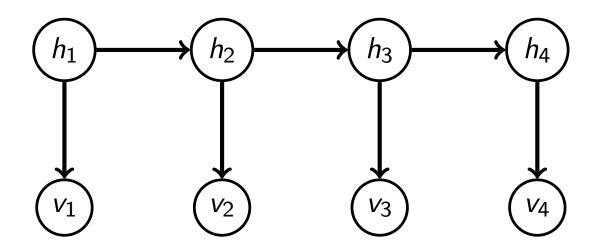
- 1. HMM parametrisation and the learning problem
- 2. Options for learning the parameters
- 3. Learning the parameters by EM

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 - Constraints on the parameters
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Hidden Markov model

Specified by

DAG (representing the independence assumptions)



- ▶ Transition distribution $p(h_i|h_{i-1})$
- ▶ Emission distribution $p(v_i|h_i)$
- ▶ Initial state distribution $p(h_1)$

The classical inference problems

- Classical inference problems:
 - ▶ Filtering: $p(h_t|v_{1:t})$
 - ▶ Smoothing: $p(h_t|v_{1:u})$ where t < u
 - ▶ Prediction: $p(h_t|v_{1:u})$ and/or $p(v_t|v_{1:u})$ where t > u
 - Most likely hidden path (Viterbi alignment): $\underset{h_{1:t}}{\operatorname{argmax}} p(h_{1:t}|v_{1:t})$
- Inference problems can be solved by message passing.
- ► Requires that the transition, emission, and initial state distributions are known.

Learning problem

▶ Data: $\mathcal{D} = \{\mathcal{D}_1, \dots, \mathcal{D}_n\}$, where each \mathcal{D}_j is a sequence of visibles of length d, i.e.

$$\mathcal{D}_j = (v_1^{(j)}, \dots, v_d^{(j)})$$

- Assumptions:
 - ▶ All variables are discrete: $h_i \in \{1, ..., K\}$, $v_i \in \{1, ..., M\}$.
 - Stationarity
- Parametrisation:
 - ► Transition distribution is parametrised by the matrix A

$$p(h_i = k | h_{i-1} = k'; \mathbf{A}) = A_{k,k'}$$

► Emission distribution is parametrised by the matrix **B**

$$p(v_i = m | h_i = k; \mathbf{B}) = B_{m,k}$$

Initial state distribution is parametrised by the vector a

$$p(h_1 = k; \mathbf{a}) = a_k$$

▶ Task: Use the data \mathcal{D} to learn \mathbf{A} , \mathbf{B} , and \mathbf{a}

Learning problem

► Since **A**, **B**, and **a** represent (conditional) distributions, the parameters are constrained to be non-negative and to satisfy

$$\sum_{k=1}^{K} p(h_i = k | h_{i-1} = k') = \sum_{k=1}^{K} A_{k,k'} = 1 \qquad \text{for all } k'$$

$$\sum_{m=1}^{M} p(v_i = m | h_i = k) = \sum_{m=1}^{M} B_{m,k} = 1 \qquad \text{for all } k$$

$$\sum_{k=1}^{K} p(h_1 = k) = \sum_{k=1}^{K} a_k = 1$$

- Note: Much of what follows holds more generally for HMMs and does not use the stationarity assumption or that the h_i and v_i are discrete random variables.
- ightharpoonup The parameters together will be denoted by θ .

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Options for learning the parameters

- ▶ The model $p(\mathbf{h}, \mathbf{v}; \boldsymbol{\theta})$ is normalised but we have unobserved variables.
- Option 1: Simple gradient ascent on the log-likelihood

$$m{ heta}_{\mathsf{new}} = m{ heta}_{\mathsf{old}} + \epsilon \sum_{j=1}^n \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_j;m{ heta}_{\mathsf{old}})} \left[
abla_{m{ heta}} \log p(\mathbf{h}, \mathcal{D}_j; m{ heta}) igg|_{m{ heta}_{\mathsf{old}}}
ight]$$

see slides Intractable Likelihood Functions

Option 2: EM algorithm

$$m{ heta}_{\mathsf{new}} = rgmax_{m{ heta}} \sum_{j=1}^n \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_j;m{ heta}_{\mathsf{old}})} \left[\log p(\mathbf{h}, \mathcal{D}_j; m{ heta})
ight]$$

see slides Variational Inference and Learning

 For HMMs, both are possible (necessary posteriors can be computed with sum-product message passing)

Options for learning the parameters

Option 1:
$$\theta_{\text{new}} = \theta_{\text{old}} + \epsilon \sum_{j=1}^{n} \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j};\theta_{\text{old}})} \left[\nabla_{\theta} \log p(\mathbf{h}, \mathcal{D}_{j}; \theta) \Big|_{\theta_{\text{old}}} \right]$$
Option 2: $\theta_{\text{new}} = \operatorname{argmax}_{\theta} \sum_{j=1}^{n} \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j};\theta_{\text{old}})} \left[\log p(\mathbf{h}, \mathcal{D}_{j}; \theta) \right]$

Similarities:

- Both require computation of the posterior expectation.
- In opt 2, assume the "M" step is performed by gradient ascent,

$$oldsymbol{ heta}' = oldsymbol{ heta} + \epsilon \sum_{j=1}^n \mathbb{E}_{p(oldsymbol{\mathsf{h}}|\mathcal{D}_j;oldsymbol{ heta}_\mathsf{old})} \left[
abla_{oldsymbol{ heta}} \log p(oldsymbol{\mathsf{h}}, \mathcal{D}_j;oldsymbol{ heta})
ight]$$

where θ is initialised with θ_{old} , and the final θ' gives θ_{new} . If only one gradient step is taken, option 2 becomes option 1.

Differences:

- ▶ Unlike option 2, option 1 requires re-computation of the posterior after each ϵ update of θ , which may be costly.
- ▶ In some cases (including HMMs), the "M"/argmax step can be performed analytically in closed form.

- 1. HMM parametrisation and the learning problem
- 2. Options for learning the parameters
- 3. Learning the parameters by EM
 - E-step
 - M-step
 - EM (Baum-Welch) algorithm

The EM objective function

▶ Denote the objective in the EM algorithm by $J(\theta, \theta_{\text{old}})$,

$$J(oldsymbol{ heta}, oldsymbol{ heta}_{\mathsf{old}}) = \sum_{j=1}^n \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_j;oldsymbol{ heta}_{\mathsf{old}})} \left[\log p(\mathbf{h}, \mathcal{D}_j; oldsymbol{ heta})
ight]$$

- Expected log-likelihood after filling-in the missing data
- We show next that for the HMM model in general, the full posteriors $p(\mathbf{h}|\mathcal{D}_i; \theta_{\text{old}})$ are not needed but just

$$p(h_i, h_{i-1} \mid \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}})$$
 $p(h_i \mid \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}}).$

They can be obtained by the alpha-beta recursion (sum-product algorithm).

Posteriors need to be computed for each observed sequence \mathcal{D}_j , and need to be re-computed after updating θ .

The EM objective function

The HMM model factorises as

$$p(\mathbf{h}, \mathbf{v}; \boldsymbol{\theta}) = p(h_1; \mathbf{a})p(v_1|h_1; \mathbf{B})\prod_{i=2}^d p(h_i|h_{i-1}; \mathbf{A})p(v_i|h_i; \mathbf{B})$$

▶ For sequence \mathcal{D}_i , we have

$$\log p(\mathbf{h}, \mathcal{D}_j; \boldsymbol{\theta}) = \log p(h_1; \mathbf{a}) + \log p(v_1^{(j)}|h_1; \mathbf{B}) + \sum_{i=2}^d \log p(h_i|h_{i-1}; \mathbf{A}) + \log p(v_i^{(j)}|h_i; \mathbf{B})$$

Since

$$\begin{split} \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(h_{1};\mathbf{a})\right] &= \mathbb{E}_{p(h_{1}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(h_{1};\mathbf{a})\right] \\ \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(h_{i}|h_{i-1};\mathbf{A})\right] &= \mathbb{E}_{p(h_{i},h_{i-1}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(h_{i}|h_{i-1};\mathbf{A})\right] \\ \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(v_{i}^{(j)}|h_{i};\mathbf{B})\right] &= \mathbb{E}_{p(h_{i}|\mathcal{D}_{j};\boldsymbol{\theta}_{\text{old}})}\left[\log p(v_{i}^{(j)}|h_{i};\mathbf{B})\right] \end{split}$$

we do not need the full posterior but only the marginal posteriors and the joint of the neighbouring variables.

The EM objective function

With the factorisation (independencies) in the HMM model, the objective function thus becomes

$$\begin{split} J(\theta, \theta_{\text{old}}) &= \sum_{j=1}^{n} \mathbb{E}_{p(\mathbf{h}|\mathcal{D}_{j}; \theta_{\text{old}})} \left[\log p(\mathbf{h}, \mathcal{D}_{j}; \theta) \right] \\ &= \sum_{j=1}^{n} \mathbb{E}_{p(h_{1}|\mathcal{D}_{j}; \theta_{\text{old}})} \left[\log p(h_{1}; \mathbf{a}) \right] + \\ &\sum_{j=1}^{n} \sum_{i=2}^{d} \mathbb{E}_{p(h_{i}, h_{i-1}|\mathcal{D}_{j}; \theta_{\text{old}})} \left[\log p(h_{i}|h_{i-1}; \mathbf{A}) \right] + \\ &\sum_{j=1}^{n} \sum_{i=1}^{d} \mathbb{E}_{p(h_{i}|\mathcal{D}_{j}; \theta_{\text{old}})} \left[\log p(v_{i}^{(j)}|h_{i}; \mathbf{B}) \right] \end{split}$$

In the derivation so far we have not yet used the assumed parametrisation of the model. We insert these assumptions next.

The term for the initial state distribution

We have assumed that

$$p(h_1 = k; \mathbf{a}) = a_k \qquad k = 1, \dots, K$$

which we can write as

$$p(h_1;\mathbf{a})=\prod_k a_k^{\mathbb{I}(h_1=k)}$$

(like for the Bernoulli model, see slides Basics of Model-Based Learning)

The log pmf is thus

$$\log p(h_1; \mathbf{a}) = \sum_k \mathbb{1}(h_1 = k) \log a_k$$

Hence

$$\mathbb{E}_{p(h_1|\mathcal{D}_j;\boldsymbol{\theta}_{\text{old}})}\left[\log p(h_1;\mathbf{a})\right] = \sum_{k} \mathbb{E}_{p(h_1|\mathcal{D}_j;\boldsymbol{\theta}_{\text{old}})}\left[\mathbb{1}(h_1=k)\right]\log a_k$$
$$= \sum_{k} p(h_1=k|\mathcal{D}_j;\boldsymbol{\theta}_{\text{old}})\log a_k$$

The term for the transition distribution

We have assumed that

$$p(h_i = k | h_{i-1} = k'; \mathbf{A}) = A_{k,k'}$$
 $k, k' = 1, ... K$

which we can write as

$$p(h_i|h_{i-1};\mathbf{A}) = \prod_{k,k'} A_{k,k'}^{\mathbb{1}(h_i=k,h_{i-1}=k')}$$

(see slides Basics of Model-Based Learning)

Further:

$$\log p(h_i|h_{i-1};\mathbf{A}) = \sum_{k,k'} \mathbb{1}(h_i = k, h_{i-1} = k') \log A_{k,k'}$$

► Hence $\mathbb{E}_{p(h_i,h_{i-1}|\mathcal{D}_i;\theta_{\text{old}})}[\log p(h_i|h_{i-1};\mathbf{A})]$ equals

$$\sum_{k,k'} \mathbb{E}_{p(h_i,h_{i-1}|\mathcal{D}_j;\theta_{\text{old}})} \left[\mathbb{1}(h_i = k, h_{i-1} = k') \right] \log A_{k,k'}$$

$$= \sum_{k,k'} p(h_i = k, h_{i-1} = k' | \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}}) \log A_{k,k'}$$

The term for the emission distribution

We can do the same for the emission distribution.

With

$$p(v_i|h_i; \mathbf{B}) = \prod_{m,k} B_{m,k}^{\mathbb{1}(v_i=m,h_i=k)} = \prod_{m,k} B_{m,k}^{\mathbb{1}(v_i=m)\mathbb{1}(h_i=k)}$$

we have

$$\mathbb{E}_{p(h_i|\mathcal{D}_j;\theta_{\text{old}})}\left[\log p(v_i^{(j)}|h_i;\mathbf{B})\right] = \sum_{m,k} \mathbb{1}(v_i^{(j)} = m)p(h_i = k|\mathcal{D}_j,\theta_{\text{old}})\log B_{m,k}$$

E-step for discrete-valued HMM

Putting all together, we obtain the EM objective function for the HMM with discrete visibles and hiddens.

$$J(\theta, \theta_{\text{old}}) = \sum_{j=1}^{n} \sum_{k} p(h_1 = k | \mathcal{D}_j; \theta_{\text{old}}) \log a_k +$$

$$\sum_{j=1}^{n} \sum_{i=2}^{d} \sum_{k,k'} p(h_i = k, h_{i-1} = k' | \mathcal{D}_j; \theta_{\text{old}}) \log A_{k,k'} +$$

$$\sum_{j=1}^{n} \sum_{i=1}^{d} \sum_{m,k} \mathbb{1}(v_i^{(j)} = m) p(h_i = k | \mathcal{D}_j, \theta_{\text{old}}) \log B_{m,k}$$

- \triangleright The objectives for **a**, and the columns of **A** and **B** decouple.
- ▶ Does not decouple in separate objectives for all parameters because of the constraint that the elements of **a** have to sum to one, and that the columns of **A** and **B** have to sum to one.

M-step

- We discuss the details for the maximisation with respect to a.
 The other cases are done equivalently.
- Optimisation problem:

$$\max_{\mathbf{a}} \sum_{j=1}^n \sum_k p(h_1 = k | \mathcal{D}_j; oldsymbol{ heta}_{\mathsf{old}}) \log a_k$$
 subject to $a_k \geq 0$ $\sum_k a_k = 1$

- ▶ The non-negativity constraint could be handled by re-parametrisation, but the constraint is here not active (the objective is not defined for $a_k \le 0$) and can be dropped.
- ► The normalisation constraint can be handled by using the methods of Lagrange multipliers (see e.g. Barber Appendix A.6).

M-step

- ▶ Lagrangian: $\sum_{j=1}^{n} \sum_{k} p(h_1 = k | \mathcal{D}_j; \theta_{\text{old}}) \log a_k \lambda(\sum_{k} a_k 1)$
- \triangleright The derivative with respect to a specific a_i is

$$\sum_{j=1}^{n} p(h_1 = i | \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}}) \frac{1}{a_i} - \lambda$$

Gives the necessary condition for optimality

$$a_i = \frac{1}{\lambda} \sum_{j=1}^n p(h_1 = i | \mathcal{D}_j; \boldsymbol{\theta}_{\mathsf{old}})$$

lacktriangle The derivative with respect to λ gives back the constraint

$$\sum_i a_i = 1$$

- ▶ Set $\lambda = \sum_{i} \sum_{j=1}^{n} p(h_1 = i | \mathcal{D}_j; \theta_{\text{old}})$ to satisfy the constraint.
- ► The Hessian of the Lagrangian is negative definite, which shows that we have found a maximum.

M-step

▶ Since $\sum_{i} p(h_1 = i | \mathcal{D}_j; \theta_{\text{old}}) = 1$, we obtain $\lambda = n$ so that

$$a_k = \frac{1}{n} \sum_{j=1}^n p(h_1 = k | \mathcal{D}_j; \boldsymbol{\theta}_{\mathsf{old}})$$

Average of all posteriors of h_1 obtained by message passing.

Equivalent calculations give

$$A_{k,k'} = \frac{\sum_{j=1}^{n} \sum_{i=2}^{d} p(h_i = k, h_{i-1} = k' | \mathcal{D}_j; \boldsymbol{\theta}_{old})}{\sum_{k} \sum_{j=1}^{n} \sum_{i=2}^{d} p(h_i = k, h_{i-1} = k' | \mathcal{D}_j; \boldsymbol{\theta}_{old})}$$

and

$$B_{m,k} = \frac{\sum_{j=1}^{n} \sum_{i=1}^{d} \mathbb{1}(v_i^{(j)} = m) p(h_i = k | \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}})}{\sum_{m} \sum_{j=1}^{n} \sum_{i=1}^{d} \mathbb{1}(v_i^{(j)} = m) p(h_i = k | \mathcal{D}_j; \boldsymbol{\theta}_{\text{old}})}$$

Inferred posteriors obtained by message passing are averaged over different sequences \mathcal{D}_j and across each sequence (stationarity).

EM for discrete-valued HMM (Baum-Welch algorithm)

Given parameters $heta_{
m old}$

1. For each sequence \mathcal{D}_i compute the posteriors

$$p(h_i, h_{i-1} \mid \mathcal{D}_j; \boldsymbol{\theta}_{old}) \qquad p(h_i \mid \mathcal{D}_j; \boldsymbol{\theta}_{old})$$

using the alpha-beta recursion (sum-product algorithm)

2. Update the parameters

$$a_{k} = \frac{1}{n} \sum_{j=1}^{n} p(h_{1} = k | \mathcal{D}_{j}; \boldsymbol{\theta}_{\text{old}})$$

$$A_{k,k'} = \frac{\sum_{j=1}^{n} \sum_{i=2}^{d} p(h_{i} = k, h_{i-1} = k' | \mathcal{D}_{j}; \boldsymbol{\theta}_{\text{old}})}{\sum_{k} \sum_{j=1}^{n} \sum_{i=2}^{d} p(h_{i} = k, h_{i-1} = k' | \mathcal{D}_{j}; \boldsymbol{\theta}_{\text{old}})}$$

$$B_{m,k} = \frac{\sum_{j=1}^{n} \sum_{i=1}^{d} \mathbb{1}(v_{i}^{(j)} = m)p(h_{i} = k | \mathcal{D}_{j}; \boldsymbol{\theta}_{\text{old}})}{\sum_{m} \sum_{j=1}^{n} \sum_{i=1}^{d} \mathbb{1}(v_{i}^{(j)} = m)p(h_{i} = k | \mathcal{D}_{j}; \boldsymbol{\theta}_{\text{old}})}$$

Repeat step 1 and 2 using the new parameters for θ_{old} . Stop e.g. if change in parameters is less than a threshold.

Program recap

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