Bioinformatics 2 - Lecture 1A

Guido Sanguinetti

School of Informatics University of Edinburgh

January 16, 2012

Definitions

- Random variables: results of non exactly reproducible experiments
- Either intrinsically random (e.g. quantum mechanics) or the system is incompletely known, cannot be controlled precisely
- The probability p_i of an experiment taking a certain value *i* is the frequency with which that value is taken in the limit of innite experimental trials
- Alternatively, we can take probability to be our belief that a certain value will be taken

More definitions

- Let x be a random variable, the set of possible values of x is the sample space Ω
- Let x and y be two random variables, p(x = i, y = j) is the joint probability of x taking value i and y taking value j (with i and j in the respective sample spaces. Often just written p(x, y) to indicate the function (as opposed to its evaluation over the outcomes i and j)
- p(x|y) is the conditional probability, i.e. the probability of x if you know y has a certain value

Rules

- Normalisation: the sum of the probabilities of all possible experimental outcomes must be 1, $\sum_{x \in \Omega} p(x) = 1$
- Sum rule: the marginal probability p(x) is given by summing the joint p(x, y) over all possible values of y,

$$p(x) = \sum_{y \in \Omega} p(x, y)$$

- Product rule: the joint is the product of the conditional and the marginal, p(x, y) = p(x|y)p(y)
- *Bayes rule*: the posterior is the ratio of the joint and the marginal

$$p(y|x) = \frac{p(x|y)p(y)}{p(x)}$$

• *Problem!* Computing the marginal is often computationally intensive

Distributions and expectations

- A probability distribution is a rule associating a number $0 \le p(x) \le 1$ to each state $x \in \Omega$, such that $\sum_{x \in \Omega} p(x) = 1$
- For finite state space can be given by a table, in general is given by a functional form
- Probability distributions (over numerical objects) are useful to compute expectations of functions

$$\langle f \rangle = \sum_{x \in \Omega} f(x) p(x)$$

Important expectations are the mean ⟨x⟩ and variance var(x) = ⟨(x - ⟨x⟩)²⟩. For more variables, also the covariance cov(x, y) = ⟨(x - ⟨x⟩)(y - ⟨y⟩)⟩ or its scaled relative the correlation corr(x, y) = cov(x, y)/√var(x)var(y)

Computing expectations

- If you know analytically the probability distribution and can compute the sums (integrals), no problem
- If you know the distribution but cannot compute the sums (integrals), enter the magical realm of approximate inference (fun but out of scope)
- If you know nothing bur have N_S samples, then use a sample approximation
- Approximate the probability of an outcome with the *frequency* in the sample

$$\langle f(x) \rangle \simeq \sum_{x} \frac{n_x}{N_S} f(x) = \frac{1}{N_S} \sum_{i=1}^{N_S} f(x_i)$$

(prove the last equality)

Independence

• Two random variables x and y are *independent* if their joint probability factorises in terms of marginals

$$p(x,y)=p(x)p(y)$$

• Using the product rule, this is equivalent to the conditional being equal to the marginal

$$p(x,y) = p(x)p(y) \Leftrightarrow p(x|y) = p(x)$$

• Exercise: if two variables are independent, then their correlation is zero. **NOT TRUE** viceversa (no correlation does not imply independence)

Continuous states

- If the state space $\boldsymbol{\Omega}$ is continuous some of the previous definitions must be modified
- The general case is mathematically difficult; we restrict ourselves to Ω = ℝⁿ and to distributions which admit a *density*, a function

$$p: \Omega \to \mathbb{R} \quad ext{s.t.} \quad p(x) \geq 0 \forall x \quad ext{and} \quad \int_{\Omega} p(x) dx = 1$$

- It can be shown that the rules of probability distributions hold also for probability densities
- Notice that p(x) is NOT the probability of the random variable being in state x (that is always zero for bounded densities); probabilities are only defined as integrals over subsets of Ω

Basic distributions

• Discrete distribution: a random variable can take N distinct values with probability $p_i = 1, ..., N$. Formally

$$p(x=i) = \prod_{j} p_{j}^{\delta_{ij}}$$

 δ_{ij} is the Kronecker delta and the p_i s form a vector of parameters.

• Dirichlet distribution: a distribution over vectors of continuous variables (p_1, \ldots, p_N) s.t. $\sum_i p_i = 1$. Its density is given by

$$f(p_1,\ldots,p_N|\alpha_1,\ldots,\alpha_N) = \frac{1}{Z}\prod_i p_i^{\alpha_i-1}$$

Z is a normalisation constant, α s are parameters

Basic distributions

• Multivariate normal: distribution over vectors x, density

$$p\left(\mathbf{x}|\boldsymbol{\mu}, \Sigma\right) = rac{1}{\sqrt{2\pi|\Sigma|}} \exp\left[-rac{1}{2}\left(\mathbf{x}-\boldsymbol{\mu}
ight)^T \Sigma^{-1}\left(\mathbf{x}-\boldsymbol{\mu}
ight)
ight]$$

How many parameters does a multivariate normal have?

• Gamma distribution: distribution over positive real numbers, density

$$p(x|k,\theta) = \frac{x^k - 1\exp(-x/\theta)}{\theta^k \Gamma(\theta)}$$

with shape parameter k and scale parameter θ

Parameters?

- Many distributions are written as conditional probabilities *given* the parameters
- Often the values of the parameters are not known
- \bullet Given observations, we can estimate them; e.g., we pick θ by maximum likelihood

$$\hat{\theta} = \operatorname{argmax}\left[\prod p(x-i|\theta)\right]$$

- Or one could place a prior distribution over the parameters
- Posteriors are computed via Bayes theorem

Exercise: fitting a discrete distribution

• We have independent observations x_1, \ldots, x_N each taking one of D possible values, giving a likelihood

$$\mathcal{L} = \prod_{i=1}^{N} p(x_i | \mathbf{p})$$

- Compute the Maximum Likelihood estimate of **p**. What is the intuitive meaning of the result? What happens if one of the *D* values is not represented in your sample?
- Alternatively, place a Dirichlet prior with parameters α over p and compute the posterior distribution. What is the meaning of the prior parameters?

Conjugate priors

- The Bayesian way has advantages in that it quantifies uncertainty and regularizes naturally
- BUT computing the normalisation in Bayes theorem is very hard
- The case when it is possible is when the prior and the posterior are of the same form (*conjugate*)
- Example: discrete and Dirichlet (exercise before)
- Exercise: conjugate priors for the univariate normal