

# Randomness and Computation

or, “Randomized Algorithms”

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## warm-up: Birthday Paradox

30 people in a room. What is the probability they share a birthday?

- ▶ Assume everyone is equally likely to be born any day (*uniform at random*). Exclude Feb 29 for neatness.
- ▶ Generate birthdays one-at-a-time from the pool of 365 (*principle of deferred decisions*).

Probability  $p_{30diff}$  that all birthdays are *different* is

$$p_{30diff} = \prod_{i=1}^{30} \frac{365 - (i - 1)}{365} = \prod_{i=1}^{30} \left(1 - \frac{(i - 1)}{365}\right) = \prod_{j=1}^{29} \left(1 - \frac{j}{365}\right).$$

Recall that  $1 + x < e^x$  for all  $x \in \mathbb{R}$ , hence  $(1 - \frac{j}{365}) < e^{-j/365}$  for any  $j$ .

## warm-up: Birthday Paradox

Hence

$$p_{30diff} < \prod_{j=1}^{29} e^{-j/365} = \left( \prod_{j=1}^{29} e^{-j} \right)^{\frac{1}{365}} = \left( e^{-\sum_{j=1}^{29} j} \right)^{\frac{1}{365}} = \left( e^{-435} \right)^{\frac{1}{365}},$$

last step using  $\sum_{j=1}^n j = \frac{n(n+1)}{2}$ .  $(e^{-435})^{\frac{1}{365}} \sim e^{-1.19} \sim 0.3$ . So with probability of at least 0.7, two people at the party share a birthday.

More general framework:

*n birthday options, m party guests*

# warm-up: General Birthday Paradox

More general framework:

*n* birthday options, *m* party guests

Probability  $p_{all-m-diff}$  that all are *different* is

$$p_{all-m-diff} = \prod_{j=1}^m \left(1 - \frac{(j-1)}{n}\right) = \prod_{j=1}^{m-1} \left(1 - \frac{j}{n}\right).$$

Continuing,

$$p_{all-m-diff} \leq \prod_{j=1}^{m-1} e^{-j/n} = \left(\prod_{j=1}^{m-1} e^{-j}\right)^{\frac{1}{n}} = \left(e^{-\sum_{j=1}^{m-1} j}\right)^{\frac{1}{n}} = e^{-\frac{(m-1)m}{2n}},$$

approximately  $e^{-m^2/2n}$ .

Suppose we set  $m = \lfloor \sqrt{n} \rfloor$ , then  $e^{-m^2/2n}$  becomes  $\sim e^{-0.5} \sim 0.6$ .

# Balls in Bins

- ▶  $m$  balls,  $n$  bins, and balls thrown *uniformly at random* into bins (usually one at a time).
- ▶ Magic bins with no upper limit on capacity.
- ▶ Common model of random allocations and their affect on overall *load* and *load balance*, typical *distribution* in the system.
- ▶ (by the birthdays analysis) we know that for  $m = \Omega(\sqrt{n})$ , then there is some constant probability  $c > 0$  of a birthday clash (BOARD).
- ▶ “Classic” question - what does the distribution look like for  $m = n$ ? Max load? (*with high probability* results are what we want).

# Balls in Bins maximum load

## Lemma (5.1)

Let  $n$  balls be thrown independently and uniformly at random into  $n$  bins. Then for sufficiently large  $n$ , the maximum load is bounded above by  $\frac{3 \ln(n)}{\ln \ln(n)}$  with probability at least  $1 - \frac{1}{n}$ .

**Proof** The probability that bin  $i$  receives  $\geq M$  balls is at most

$$\binom{n}{M} \frac{n^{m-M}}{n^m} = \binom{n}{M} \frac{1}{n^M}.$$

Expanding  $\binom{n}{M}$ , this is

$$\frac{n \dots (n - M + 1)}{M!} \frac{1}{n^M} \leq \frac{1}{M!}.$$

To bound  $(M!)^{-1}$  note that for any  $k$ , we have  $\frac{k^k}{k!} \leq \sum_{i=0}^{\infty} \frac{k^i}{i!} = e^k$ ,  
hence  $\frac{1}{k!} \leq \left(\frac{e}{k}\right)^k$ . Or use Stirling ...

# Balls in Bins maximum load

## Proof of Lemma 5.1 cont'd.

So bin  $i$  gets  $\geq M$  balls with probability at most

$$\left(\frac{e}{M}\right)^M.$$

Set  $M =_{\text{def}} \frac{3 \ln(n)}{\ln \ln(n)}$ . Then the probability that *any* bin gets  $\geq M$  balls is (using the Union bound) at most

$$n \cdot \left(\frac{e \cdot \ln \ln(n)}{3 \ln(n)}\right)^{\frac{3 \ln(n)}{\ln \ln(n)}} \leq n \cdot \left(\frac{\ln \ln(n)}{\ln(n)}\right)^{\frac{3 \ln(n)}{\ln \ln(n)}} = e^{\ln(n)} \left(\frac{\ln \ln(n)}{\ln(n)}\right)^{\frac{3 \ln(n)}{\ln \ln(n)}}.$$

Again using properties of  $\ln$ , this expands as

$$e^{\ln(n)} \left(e^{\ln \ln \ln(n) - \ln \ln(n)}\right)^{\frac{3 \ln(n)}{\ln \ln(n)}} = e^{\ln(n)} \left(e^{-3 \ln(n) + 3 \frac{\ln(n) \ln \ln \ln(n)}{\ln \ln(n)}}\right).$$



# Balls in Bins maximum load

## Proof of Lemma 5.1 cont'd.

Grouping the  $\ln(n)$ s in the exponents, and evaluating, we have

$$e^{-2 \ln(n)} \cdot e^{3 \frac{\ln(n) \ln \ln \ln(n)}{\ln \ln(n)}} = \frac{1}{n^2} n^{3 \frac{\ln \ln \ln(n)}{\ln \ln(n)}}.$$

If we take  $n$  “sufficiently large” ( $n \geq e^{e^{e^4}}$  will do it), then  $\frac{\ln \ln \ln(n)}{\ln \ln(n)} \leq 1/3$ , hence the probability of *some* bin having  $\geq M$  balls is at most

$$\frac{1}{n}.$$

□

Can derive a matching proof to show that “with high probability” there will be a bin with  $\Omega\left(\frac{\ln(n)}{\ln \ln(n)}\right)$  balls in it. We are going to skip over this (can read in Sections 5.3 and 5.4, won't be examined)

## $\Omega(\cdot)$ bound on the maximum load (chat)

- ▶ We implicitly used the *Union Bound* in our proof Lemma 5.1, when we multiplied by  $n$  on slide 7. However, in reality, bin  $i$  has a lower chance of being “high” (say  $\Omega(\frac{\ln(n)}{\ln \ln(n)})$ ) if other bins are already “high” (the “high-bin” events are *negatively correlated*).
- ▶ This means that we can’t use the same approach as in Theorem 5.1 to prove a partner result of  $\Omega(\frac{\ln(n)}{\ln \ln(n)})$ .
- ▶ Solution is to use the fact that for the binomial distribution  $B(m, \frac{1}{n})$  for an individual bin, that as  $n \rightarrow \infty$ ,

$$\Pr[X = k] = \binom{m}{k} \left(\frac{1}{n}\right)^k \left(1 - \frac{1}{n}\right)^{m-k} \rightarrow \frac{e^{-m/n} (m/n)^k}{k!}$$

(ie, close to the probabilities for the Poisson distribution with parameter  $\mu = m/n$ )

- ▶ The Poisson’s aren’t independent but the dependance can be limited to an extra factor of  $e\sqrt{m}$  (Section 5.4).

# Average-case analysis of Bucket Sort

- ▶ Items to be sorted are natural numbers from some bounded range  $[0, 2^k)$ , some large  $k$ .
- ▶ We have a collection of empty “buckets” (extendable arrays or lists).
- ▶ Each bucket has an “index” used to access it.
- ▶ We have some value  $m$ , the “number of prefix bits” (substantially smaller than  $k$ ). We will have a bucket for each individual  $\{0, 1\}^m$ .

**Algorithm** BUCKETSORT( $a_1, \dots, a_n$ )

1. Do a linear scan of the inputs, adding  $a_i$  to the bucket matching its first  $m$  bits.
2. **for** every  $b \in \{0, 1\}^m$  **do**
3.       Sort bucket  $b$  with any  $O(n^2)$  sorting algorithm.

# Average-case analysis of Bucket Sort

Imagine that we draw the  $n$  inputs to BUCKETSORT independently and uniformly at random from  $\{0, 1\}^k$ . Hence ...

The first- $m$ -bits of the inputs are independently uniform from  $\{0, 1\}^m$ .

Each  $a_i$  has probability  $\frac{1}{2^m}$  of entering any bucket.

Bucket Sort can be seen as a “balls-in-bins” experiment.

Running time is  $\Theta(n)$  for the linear scan of 1. The *expected* running time for 2.-3. will be  $E[\sum_{b \in \{0,1\}^m} c \cdot (X_b^2)]$ , where  $X_b$  is the number of inputs landing in bucket  $b$ , and  $c > 0$  is the fixed constant of the  $O(n^2)$  algorithm.

We want to evaluate  $E[\sum_{b \in \{0,1\}^m} c \cdot (X_b^2)] = \sum_{b \in \{0,1\}^m} c \cdot E[X_b^2]$ .

We are now going to use an unexpected “trick” where we exploit the “second moment” of Binomial random variables to bound the  $E[X_b^2]$ .

# Average-case analysis of Bucket Sort

Realise each  $X_b$  is a binomial random variable  $B[n, \frac{1}{2^m}]$  with

$$E[X_b^2] = n(n-1)2^{-2m} + n2^{-m}.$$

Multiplying by  $2^m$  (for each  $b \in \{0, 1\}^m$ ), and by  $c$ , this gives expected time for 2.-3. at most

$$c \cdot (n^2 2^{-m} + n).$$

Choose  $m$  carefully to satisfy  $m \geq \lg(n)$  and we see that this ensures the expected number of steps for 2.-3. is at most  $2 \cdot c \cdot n$ .

# The rest of the course

Lect 11 Random Graphs and Hamilton cycles (Section 5.6)

Lects 12-13 The Probabilistic method, derandomization via Conditional expectation (bit more than half Chapter 6)

Will hold a “tutorial” in the lecture slot for Friday 10th March (we will cover questions about Coursework 1, and “end of printout” questions between now and then)

Lects 14-15 Markov chain basics (first half Chapter 7)

Lects 16-17 The Monte Carlo method (some of Chapter 9)

Lects 18-20 Mixing time bounds for Markov chains (Chapter 11)

I will hold the second “tutorial” in the Lecture slot of Friday 7th April (our final meeting).

# References and Exercises

- ▶ Sections 5.1, 5.2 of “Probability and Computing”. And if you are interested in the  $\Omega$  bound for the  $\Theta\left(\frac{\ln(n)}{\ln \ln(n)}\right)$  result, read Sections 5.3 and 5.4 also.
- ▶ Section 5.5 on Hashing is worth a read and has none of the Poisson stuff (I’m skipping it because of time limitations).

## Exercises

- ▶ Exercise 5.3 (balls in bins when  $m = c \cdot \sqrt{n}$ ).
- ▶ Exercise 5.10 (sequences of empty bins; this is a bit more tricky)