

Probabilistic Modelling and Reasoning, Tutorial Answer Sheet 3 (for Week 5)

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Prob. Sheet. ID: 11922.1.2

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1.

$$P(n_h, n_t | \theta) = \theta^{n_h} (1 - \theta)^{n_t} \quad (\text{Binomial likelihood})$$

$$\hat{\theta}_{\text{ML}} = \frac{n_h}{n_h + n_t} \quad (\text{Maximum likelihood solution, see last week's tutorial})$$

$$p(\theta) \propto \theta^{\alpha_h - 1} (1 - \theta)^{\alpha_t - 1} \quad (\text{Beta prior})$$

$$\text{E}[\theta] = \frac{\alpha_h}{\alpha_h + \alpha_t} \quad (\text{Prior mean, property of beta distribution})$$

$$p(\theta | n_h, n_t) \propto P(n_h, n_t | \theta) p(\theta) \propto \theta^{\alpha_h + n_h - 1} (1 - \theta)^{\alpha_t + n_t - 1} \quad (\text{Posterior is also a beta distribution})$$

$$\text{E}[\theta | \mathcal{D}] = \frac{\alpha_h + n_h}{\alpha_h + \alpha_t + n_h + n_t} \quad (\text{Posterior mean, property of beta distribution})$$

$$= \frac{\alpha_h}{\alpha + n} + \frac{n_h}{\alpha + n} \quad (\alpha = \alpha_h + \alpha_t, n = n_h + n_t)$$

$$= \frac{\alpha}{\alpha + n} \frac{\alpha_h}{\alpha} + \frac{n}{\alpha + n} \frac{n_h}{n}$$

$$= \lambda \text{E}[\theta] + (1 - \lambda) \hat{\theta}_{\text{ML}} \quad (\lambda = \alpha / (\alpha + n))$$

2.

$$\frac{d \log p}{d\theta} = \frac{\alpha - 1}{\theta} - \frac{\beta - 1}{1 - \theta}$$

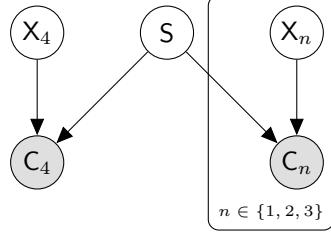
$$\frac{\theta^*}{\alpha - 1} = \frac{1 - \theta^*}{\beta - 1} \quad \text{at maximum}$$

$$\theta^* = \frac{\alpha - 1}{\alpha + \beta - 2} \quad \text{is the value of } \theta \text{ at maximum}$$

For same positions of mean and mode

$$\begin{aligned} \frac{\alpha - 1}{\alpha + \beta - 2} &= \frac{\alpha}{\alpha + \beta} \\ \implies \alpha &= \beta \end{aligned}$$

3. The graphical model and conditional probabilities are given by.



$$\begin{aligned}\mathbb{P}[X_n = x] &= \begin{cases} \alpha & : x = 1 \\ 1 - \alpha & : x = 2 \end{cases} \\ \mathbb{P}[S = s] &= \begin{cases} \beta & : x = 1 \\ 1 - \beta & : x = 2 \end{cases} \\ \mathbb{P}[C_n = c | X_n = x, S = s] &= \begin{cases} 1 - \lambda & : c = 1, x = 1, s = 1 \\ \lambda & : c = 2, x = 1, s = 1 \\ \delta & : c = 1, x = 2, s = 1 \\ 1 - \delta & : c = 2, x = 2, s = 1 \\ \lambda & : c = 1, x = 1, s = 2 \\ 1 - \lambda & : c = 2, x = 1, s = 2 \\ 1 - \delta & : c = 1, x = 2, s = 2 \\ \delta & : c = 2, x = 2, s = 2 \end{cases}\end{aligned}$$

Here $\alpha = 0.3, \beta = 0.25, \gamma = 0.1, \delta = 0.01$.

Given data $\mathbf{c}_{1:3} = \{1, 1, 2\}$ probability of next crowd-sourced image label

$$\begin{aligned}\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] &= \sum_{x_4 \in \{1, 2\}} \sum_{s \in \{1, 2\}} \{\mathbb{P}[C_4 = 1, X_4 = x_4, S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\} \\ &= \sum_{x_4 \in \{1, 2\}} \sum_{s \in \{1, 2\}} \{\mathbb{P}[C_4 = 1 | X_4 = x_4, S = s, \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] \mathbb{P}[X_4 = x_4 | S = s, \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] \mathbb{P}[S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\} \\ &= \sum_{s \in \{1, 2\}} \left(\sum_{x_4 \in \{1, 2\}} \{\mathbb{P}[C_4 = 1 | X_4 = x_4, S = s] \mathbb{P}[X_4 = x_4]\} \mathbb{P}[S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] \right)\end{aligned}$$

Probability of supervisor identity given data

$$\begin{aligned}\mathbb{P}[S = s | \mathbf{C}_{1:3} = \{1, 1, 2\}] &\propto \mathbb{P}[\mathbf{C}_{1:3} = \{1, 1, 2\} | S = s] \mathbb{P}[S = s] \\ &= \prod_{n \in \{1, 2, 3\}} \{\mathbb{P}[C_n = c_n | S = s]\} \mathbb{P}[S = s] \\ &= \prod_{n \in \{1, 2, 3\}} \left\{ \sum_{x_n \in \{1, 2\}} (\mathbb{P}[C_n = c_n | X_n = x_n, S = s] \mathbb{P}[X_n = x_n]) \right\} \mathbb{P}[S = s]\end{aligned}$$

Substituting in values gives

$$\begin{aligned}
\mathbb{P}[S = 1 | \mathbf{C}_{1:3} = \{1, 1, 2\}] &\propto ((1 - \lambda)\alpha + \delta(1 - \alpha))^2 (\lambda\alpha + \delta(1 - \alpha)) \beta \\
&= 0.277^2 \times 0.723 \times 0.75 = 0.0416 \\
\mathbb{P}[S = 2 | \mathbf{C}_{1:3} = \{1, 1, 2\}] &\propto (\lambda\alpha + \delta(1 - \alpha))^2 ((1 - \lambda)\alpha + \delta(1 - \alpha)) (1 - \beta) \\
&= 0.723^2 \times 0.277 \times 0.25 = 0.0362 \\
\therefore \mathbb{P}[S = 1 | \mathbf{C}_{1:3} = \{1, 1, 2\}] &= \frac{0.0416}{0.0416 + 0.0362} = 0.535 = \epsilon \\
\text{and } \mathbb{P}[S = 2 | \mathbf{C}_{1:3} = \{1, 1, 2\}] &= \frac{0.0362}{0.0416 + 0.0362} = 0.465 = 1 - \epsilon
\end{aligned}$$

Substituting this into above gives

$$\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] = [(1 - \gamma)\alpha + \delta + \delta(1 - \alpha)] \epsilon + [\gamma\alpha + (1 - \delta)(1 - \alpha)] (1 - \epsilon) = 0.485 = \zeta$$

For final part, the probability of the true label being a car given the now observed crowd-sourced label $C_4 = 1$

$$\begin{aligned}
\mathbb{P}[X_4 = 1 | C_4 = 1, \mathbf{C}_{1:3} = \mathbf{c}_{1:3}] &= \frac{\mathbb{P}[X_4 = 1, C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]}{\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]} \\
&= \frac{\sum_{s \in \{1, 2\}} \{\mathbb{P}[X_4 = 1, C_4 = 1, S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\}}{\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]} \\
&= \frac{\sum_{s \in \{1, 2\}} \{\mathbb{P}[C_4 = 1 | X_4 = 1, S = s, \cancel{\mathbf{C}_{1:3} = \mathbf{c}_{1:3}}] \mathbb{P}[X_4 = 1, S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\}}{\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]} \\
&= \frac{\sum_{s \in \{1, 2\}} \{\mathbb{P}[C_4 = 1 | X_4 = 1, S = s] \mathbb{P}[X_4 = 1 | S = s, \cancel{\mathbf{C}_{1:3} = \mathbf{c}_{1:3}}] \mathbb{P}[S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\}}{\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]} \\
&= \frac{\mathbb{P}[X_4 = 1] \sum_{s \in \{1, 2\}} \{\mathbb{P}[C_4 = 1 | X_4 = 1, S = s] \mathbb{P}[S = s | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]\}}{\mathbb{P}[C_4 = 1 | \mathbf{C}_{1:3} = \mathbf{c}_{1:3}]} \\
&= \frac{\alpha \{(1 - \gamma)\epsilon + \gamma(1 - \epsilon)\}}{\zeta} = 0.327
\end{aligned}$$