# Learning from Data: Density Estimation -Gaussian Distribution

#### Amos Storkey, School of Informatics

Semester 1

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#### Worked Example

- ► 3 class data: e.g. 1,2,1,3,1,1,2,1,2,3,1,1,1.
- Represent in 1 of *m*:  $c_i = 1$  iff (if and only if) class is *i*.
- Write out likelihood of one datum. Use  $\theta_i = P(c_i = 1)$

$$P(c_i = 1 | \Theta) = \prod_i \theta_i^{c_i} = \theta_1^{c_1} \theta_2^{c_2} \theta_3^{c_3}$$

• (Note: 
$$\theta_1^{c_1} \theta_2^{c_2} \theta_3^{c_3} = \theta_2$$
 iff  $c_2 = 1$ ) etc.

Now for all the data D:

$$\mathcal{P}(\mathcal{D}|\Theta) = \prod_{\mu} \prod_{i} heta_{i}^{c_{i}^{\mu}} = heta_{1}^{N_{1}} heta_{2}^{N_{2}} heta_{3}^{N_{3}}$$

Take logs

$$\log P(D|\Theta) = \sum_{\mu} \sum_{i} c_{i}^{\mu} \log \theta_{i} = \sum_{i} N_{i} \log \theta_{i}$$

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### Worked Example

Take logs

$$\log \mathcal{P}(\mathcal{D}|\Theta) = \sum_{\mu} \sum_{i} c^{\mu}_{i} \log heta_{i} = \sum_{i} \mathcal{N}_{i} \log heta_{i}$$

Need to optimise subject to condition ∑<sub>i</sub> θ<sub>i</sub> = 1. Add on Lagrange multiplier term λ(∑<sub>i</sub> θ<sub>i</sub> − 1), and differentiate wrt θ<sub>k</sub> using

$$rac{\partial}{\partial heta_k} [\log \mathcal{P}(\mathcal{D}|\Theta) + \lambda (\sum_i heta_i - 1)] = rac{\mathcal{N}_k}{ heta_k} + \lambda$$

Set derivative to zero to get θ<sub>k</sub> = −N<sub>k</sub>/λ. Substitute into constraint: Σ<sub>i</sub> θ<sub>i</sub> = 1 to get λ = −Σ<sub>k</sub> N<sub>k</sub>. Final answer:

$$\theta_k = \frac{N_k}{\sum_k N_k}$$

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- This lecture we will be focusing on continuous quantities.
- The most common (and most easily analysed) distribution for continuous quantities is the Gaussian distribution.
- Gaussian distribution is often a reasonable model for many quantities due to various central limit theorems.
- Gaussian is sometimes called a normal distribution.

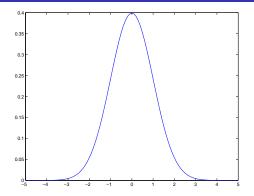
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The one dimensional Gaussian distribution is given by

$$P(x|\mu, \sigma^2) = N(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- $\mu$  is the *mean* of the Gaussian and  $\sigma^2$  is the *variance*.
- If μ = 0 and σ<sup>2</sup> = 1 then N(x; μ, σ<sup>2</sup>) is called a *standard* Gaussian.

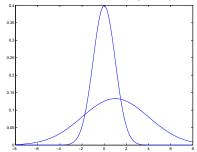
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- This is a standard one dimensional Gaussian distribution.
- All Gaussians have the same shape subject to scaling and displacement.
- If x is distributed N(x; μ, σ<sup>2</sup>), then y = (x − μ)/σ is distributed N(y; 0, 1).

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- Remember all distributions must integrate to one. The  $\sqrt{2\pi\sigma^2}$  is called a normalisation constant it ensures this is the case.
- Hence tighter Gaussians have higher peaks:



- >  $X_i$  mean 0, variance  $\Sigma$ , not necessarily Gaussian.
- ► X<sub>i</sub> subject to various conditions (e.g. IID).

$$\frac{1}{\sqrt{N}}\sum_{i=1}^N X_i \sim N(0, \Sigma)$$

asymptotically as  $N \rightarrow \infty$ .

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- Suppose we have data  $\{x_i, i = 1, 2, ..., n\}$ .
- Suppose we presume the data was generated from a Gaussian with mean μ and variance σ<sup>2</sup>. Call this the model.
- Then the log probability of the data given the model is

$$\log \prod_{i} P(x_{i}|\mu,\sigma^{2}) = -\frac{1}{2} \sum_{i} \frac{(x_{i}-\mu)^{2}}{\sigma^{2}} - \frac{N}{2} \log(2\pi\sigma^{2})$$

Steps left as exercise: hint  $\log \prod = \sum \log$ 

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### Maximum Likelihood Estimation

• Maximum likelihood: Set  $\gamma = 1/\sigma^2$  Take derivatives

$$\log P(X|\mu,\gamma) = -\frac{1}{2} \sum_{i} \gamma(x_{i}-\mu)^{2} - \frac{N}{2} \log(2\pi) + \frac{N}{2} \log \gamma$$
$$\frac{\partial \log P(X|\mu,\gamma)}{\partial \mu} = \gamma \sum_{i} (x_{i}-\mu)$$
$$\frac{\partial \log P(X|\mu,\gamma)}{\partial \gamma} = -\frac{1}{2} \sum_{i} (x_{i}-\mu)^{2} + \frac{N}{2\gamma}$$

- Hence  $\mu = (1/N) \sum_{i} x_{i}$  and  $\sigma^{2} = (1/N) \sum_{i} (x_{i} \mu)^{2}$ .
- (Maximum likelihood estimate of  $\sigma^2$  is biased.)

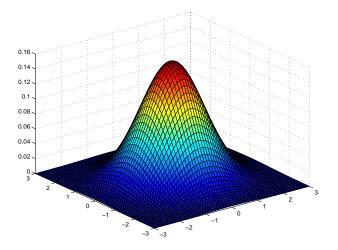
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The vector x is multivariate Gaussian if for mean μ and covariance matrix Σ, it is distributed according to

$$P(\mathbf{x}|\boldsymbol{\mu},\boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{d/2}|\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T\boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right)$$

- The univariate Gaussian is a special case of this.
- Σ is called a covariance matrix. It says how much attributes co-vary. More later.

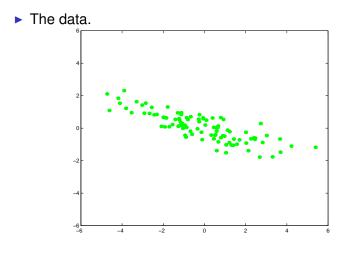
#### Multivariate Gaussian: Picture



- The Maximum Likelihood estimate can be found in the same way.
- $\boldsymbol{\mu} = (1/N) \sum_{i=1}^{N} \mathbf{x}_i$   $\boldsymbol{\Sigma} = (1/N) \sum_{i=1}^{N} (\mathbf{x}_i \boldsymbol{\mu}) (\mathbf{x}_i \boldsymbol{\mu})^T$

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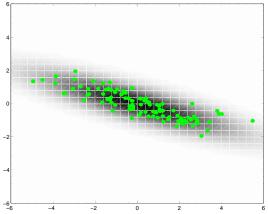
Example



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#### The data. The maximum likelihood fit.



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- Have real valued multivariate data, along with class label for each point.
- Want to predict the value of the class label given some new point.
- Presume that if we take all the points with a particular label, then we believe they were sampled from a Gaussian.
- How should we predict the class at a new point?

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## Class conditional classification

- Learning: Fit Gaussian to data in each class (class conditional fitting). Gives P(position|class)
- Find estimate for probability of each class (see last lecture)
  P(class)
- Inference: Given a new position, we can ask "What is the probability of this point being generated by each of the Gaussians?"
- Pick the largest (just like maximum likelihood)
- Better still give probability using Bayes rule

 $P(\text{class}|\text{position}) \propto P(\text{position}|\text{class})P(\text{class})$ 

Then can get ratio

P(class = 1|position)/P(class = 0|position).

Decision boundary for two classes is where this ratio is one.

- Gaussian
- Maximum Likelihood fitting of a Gaussian
- Multivariate Gaussian and covariances again.
- Maximum Likelihood fitting.
- Class conditional classification using Gaussians.

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