# Ab initio prediction of protein interaction

### **Dirk Husmeier**

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## Pathways and systems biology





## Protein-Protein



Metabolic Pathways

### Molecular Interaction Networks

Apoptosis



Signaling Pathways Lac Operon



Gene Regulation



SH3 domain protein interaction network in S. cerevisiae; from Tong et al. (2002)

Experimental high-throughput techniques

Yeast two-hybrid Phage display



Fig. 1. Interaction detection by yeast two-hybrid assay. (a) Activation of reporter gene by transcriptional activator; (b) Activation of reporter gene by reconstituted transcriptional activator.

From See-Kiong Ng and Soon-Heng Tan, J. Bioinf. Comp. Bio. (2004)



Fig. 2. Schematic diagrams of (a) a phage; and (b) interaction detection by phage display. From See-Kiong Ng and Soon-Heng Tan, J. Bioinf. Comp. Bio. (2004)

Tong et al. (2002), Science 295, 321-324. SH3 domain proteins in *Saccharomyces cerevisiae*. **Yeast two-hybrid** interaction network 285 interactions between 28 SH3 proteins and 143 binding peptides Phage display interaction network 394 interactions between 28 SH3 proteins and 178 binding peptides



- High-throughput experiments (yeast two-hybrid, phage display) are expensive and intrinsically noisy.
- It would be desirable to more specifically target or partially bypass them with complementary *in silico* approaches.

- High-throughput experiments (yeast two-hybrid, phage display) are expensive and intrinsically noisy.
- It would be desirable to more specifically target or partially bypass them with complementary *in silico* approaches.
- Objective: develop a probabilistic model to predict protein-protein interactions from sequence data.
- Method: We want to capture the way protein recognition modules recognise and bind to peptide ligands that contain a specific binding motif.





Peptide recognition modules

Example: SH3 domain

			MYC3		PPPP TROPAC
		27 YPR164W 25 24 29			
	2	BP CONSERV	EDDDDDDDD KALL (SRS	EDVDVDVD S S S S S S S S S S S S S S S S	e proposition and the proposition of the propositio
VRP1	RVS167	20 YNL094W	BZZ 1	20 LAS17	SLA1
			PPPPPP CASES		TECTOLIAL
BUD14	YHL002W	a BOI2	BNI1	BNR1	YGR196W
TPAL REP ISKSPLITE			PP///////		₽ PD IDS
YHR016C		ABP1	5HO1		
P V DV PR Resault	BLATE STOR		BAYESSAN		



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## TCGAATTCTATA GCCAC

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TCGAA TTCTATA G C C A C



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## Position Specific Scoring Matrix (PSSM)

Search for a motif of length W in binding sequences.

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Position Specific Scoring Matrix (PSSM) Search for a motif of length W in binding sequences.  $W \times 4$  matrix  $\psi_k(l)$ : Probability that the nucleotide in the kth position,  $k \in [1, ..., W]$ , is an  $l \in \{A, C, G, T\}$ . Background model for non-binding sequences 4-dim vector  $\theta_0(l)$ :

Probability of nucleotide l; this distribution is position-independent.

## Sequence $S_1, S_2, \ldots, S_N$

Sequence  $S_1, S_2, \dots, S_N$ Non-binding sequence: R=0  $P(S_1, S_2, \dots, S_N | R = 0) = \prod_{t=1}^N \theta_0(S_t)$  Sequence  $S_1, S_2, \dots, S_N$ Non-binding sequence: R=0  $P(S_1, S_2, \dots, S_N | R = 0) = \prod_{t=1}^N \theta_0(S_t)$ 

Binding sequence: R=1, motif starting at position m+1

$$P(S_1, S_2, \dots, S_N | R = 1, start = m + 1)$$

$$= \prod_{t=1}^{m} \theta_0(S_t) \prod_{k=1}^{W} \psi_k(S_{m+k}) \prod_{t=m+W+1}^{N} \theta_0(S_t)$$

$$= \prod_{t=1}^{N} \theta_0(S_t) \prod_{k=1}^{W} \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}$$

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Binding sequence: R=1, motif starting anywhere

$$P(S_1, S_2, \dots, S_N | R = 1)$$

$$= \sum_{m=0}^{N-W} P(start = m+1) P(S_1, S_2, \dots, S_N | R = 1, start = m+1)$$

$$= \prod_{t=1}^{N} \theta_0(S_t) \frac{1}{N-W+1} \sum_{m=0}^{N-W} \prod_{k=1}^{W} \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}$$

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## Gibbs sampling











Posterior probability of a binding location, given the parameters

$$P(start = m + 1 | S_1, S_2, \dots, S_N, Parameters) = \frac{P(S_1, S_2, \dots, S_N | start = m + 1)P(start = m + 1)}{P(S_1, S_2, \dots, S_N)}$$

Motif starting at position m+1

$$P(S_1, S_2, \dots, S_N | start = m + 1) = \prod_{t=1}^N \theta_0(S_t) \prod_{k=1}^W \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}$$

Motif starting anywhere

$$P(S_1, S_2, \dots, S_N) = \sum_{m=0}^{N-W} P(start = m+1)P(S_1, S_2, \dots, S_N | start = m+1)$$


Posterior probability of the parameters, given the binding locations

Sufficient statistics: Count matrix  $C_{k,l}$ 

 $C_{k,l}$ : Number of times amino acid l appears in position k.

$$P(D, binding \ locations | parameters) = \prod_{k=1}^{W} \prod_{l=1}^{20} \psi_{k,l}^{C_{k,l}}$$

Conjugate prior distribution: Dirichlet

$$P(parameters) \propto \prod_{l=1}^{20} \psi_{k,l}^{\alpha_l - 1}$$

Posterior distribution

 $P(parameters|D, binding \ locations) \circ$ 

$$\prod_{k=1}^{W} \prod_{l=1}^{20} \psi_{k,l}^{\mathbf{C}_{k,l} + \alpha_l - 1}$$



Counts = C

A	0	0	0	0	0
С	0	0	0	0	0
G	0	0	0	0	0
Т	0	0	0	0	0
	1	2	3	4	5













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# Sufficient statistics: $C_{d,s}$ $C_{d,s,k,l} = \delta(sequence_{s,a_{d,s}+k} = l)$

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Sequence s

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### H P K W S P L P P W H K

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Sufficient statistics: 
$$C_{d,s}$$
  
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$$C_{d,s,k,l} = \delta(sequence_{s,a_{d,s}+k} = l)$$

In words:  $C_{d,s,k,l}$  is 1 if the *k*th position of the binding motif in sequence *s* that binds to PRM domain *d* is amino acid *l*. Otherwise, it is zero.

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## Problem:

There are too few binding peptide sequences (average: 9 sequences per domain)  $\Rightarrow$  high estimation uncertainty.

#### **BIOINFORMATICS**



#### Predicting protein–peptide interactions via a network-based motif sampler

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Institute for Systems Biology, 1441 North 34th street, Seattle, WA 98103-8904, USA

Received on January 15, 2004; accepted on March 1, 2004

### Modify the count matrix $C_{d,s}$ , using the network topology.



 $\mathbf{ ilde{C}}_{d,s}$ 



$$ilde{\mathbf{C}}_{d,s} = \sum_{s} \varepsilon_{d,s} \mathbf{C}_{d,s}$$



 $\tilde{\mathbf{C}}_{d,s} = \sum_{s} \varepsilon_{d,s} \mathbf{C}_{d,s} + \lambda_1 \sum_{d} \varepsilon_{d,s} \mathbf{C}_{d,s}$ 



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SH3 domain protein interaction network in S. cerevisiae; from Tong et al. (2002)

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Reiss & Schwikowski (2004)

 $\tilde{\mathbf{C}}_{d,s} = \sum_{s} \varepsilon_{d,s} \mathbf{C}_{d,s} + \lambda_{1} \sum_{d} \varepsilon_{d,s} \mathbf{C}_{d,s} + \lambda_{2} \sum_{s} \sum_{d} \varepsilon_{d,s} \mathbf{C}_{d,s}$ 

### Reiss & Schwikowski (2004)

## $\tilde{\mathbf{C}}_{d,s} = \sum_{s} \varepsilon_{d,s} \mathbf{C}_{d,s} + \lambda_{1} \sum_{d} \varepsilon_{d,s} \mathbf{C}_{d,s} + \lambda_{2} \sum_{s} \sum_{d} \varepsilon_{d,s} \mathbf{C}_{d,s}$

Heuristic modification to make the model more discriminative: Give higher probability to sites that are distinct from non-binding motifs. New tuning parameter  $\lambda_3$ 

Sequence analysis

#### A regularized discriminative model for the prediction of protein–peptide interactions

Wolfgang P. Lehrach<sup>1,2,\*</sup>, Dirk Husmeier<sup>2</sup> and Christopher K. I. Williams<sup>1</sup> <sup>1</sup>University of Edinburgh, Edinburgh EH1 2QL, UK and <sup>2</sup>Biomathematics and Statistics Scotland, Edinburgh EH9 3JZ, UK Received on August 4, 2005; revised on November 23, 2005; accepted on November 25, 2005

Received on August 4, 2005; revised on November 23, 2005; accepted on November 25, 2005 Advance Access publication January 5, 2006 Associate Editor: Keith A Crandall Binding sequence: R=1, motif starting at position m+1

$$P(S_1, S_2, \dots, S_N | R = 1, start = m + 1) = \prod_{t=1}^N \theta_0(S_t) \prod_{k=1}^W \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}$$
Binding sequence: R=1, motif starting anywhere

$$P(S_1, S_2, \dots, S_N | R = 1)$$

$$= \sum_{m=0}^{N-W} P(start = m+1) P(S_1, S_2, \dots, S_N | R = 1, start = m+1)$$

$$= \prod_{t=1}^{N} \theta_0(S_t) \frac{1}{N-W+1} \sum_{m=0}^{N-W} \prod_{k=1}^{W} \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}$$

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Objective: Prediction of binding activity from sequence:  $P(R = 1 | S_1, S_2, ..., S_N)$ 

#### Apply Bayes rule:

$$P(R = 1|S_1, S_2, \dots, S_N) = \frac{P(S_1, S_2, \dots, S_N | R = 1)P(R = 1)}{P(S_1, S_2, \dots, S_N)}$$
  
= 
$$\frac{P(S_1, S_2, \dots, S_N | R = 1)P(R = 1)}{P(S_1, S_2, \dots, S_N | R = 0)P(R = 0) + P(S_1, S_2, \dots, S_N | R = 1)P(R = 1)}$$
  
= 
$$\left(1 + \frac{P(R = 0)P(S_1, S_2, \dots, S_N | R = 0)}{P(R = 1)P(S_1, S_2, \dots, S_N | R = 1)}\right)^{-1}$$
  
= 
$$\left(1 + \left[\frac{P(R = 1)}{P(R = 0)}\frac{1}{(N - W + 1)}\sum_{m=0}^{N - W}\prod_{k=1}^{W}\frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}\right]^{-1}\right)^{-1}$$

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=  $\frac{P(S_1, S_2, \dots, S_N | R = 1) P(R = 1)}{P(S_1, S_2, \dots, S_N | R = 0) P(R = 0) + P(S_1, S_2, \dots, S_N | R = 1) P(R = 1)}$   
=  $\left(1 + \frac{P(R = 0) P(S_1, S_2, \dots, S_N | R = 0)}{P(R = 1) P(S_1, S_2, \dots, S_N | R = 1)}\right)^{-1}$   
=  $\left(1 + \left[\frac{P(R = 1)}{P(R = 0)} \frac{1}{(N - W + 1)} \sum_{m=0}^{N - W} \prod_{k=1}^{W} \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}\right]^{-1}\right)^{-1}$ 

### Define:

$$w_k(l) = \log \frac{\psi_k(l)}{\theta_0(l)}, \ w_0 = \frac{P(R=1)}{P(R=0)}, \ \log it(z) = \frac{1}{1 + \exp(-z)}$$

$$P(R = 1 | S_1, S_2, \dots, S_N) = \log \left( \log \left[ \frac{w_0}{N - W + 1} \sum_{m=0}^{N - W} \exp \left( \sum_{k=1}^{W} w_k(S_{t+k}) \right) \right] \right)$$

4 imes W + 1 parameters:  $w_k(l)$  ,  $w_0$ 

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$$P(R = 1 | S_1, S_2, \dots, S_N) = \log \left( \log \left[ \frac{w_0}{N - W + 1} \sum_{m=0}^{N - W} \exp \left( \sum_{k=1}^{W} w_k(S_{t+k}) \right) \right] \right)$$

4 imes W + 1 parameters:  $w_k(l)$  ,  $w_0$ 

Data D: Set of sequences  $\mathbf{x}_i$  with associated interaction indicators  $R_i \in \{0, 1\}$ 

Model predicts an interaction  $R_i$  given the sequence  $\mathbf{x}_i$ :

$$y(\mathbf{x}_i, \mathbf{w}) = P(R_i = 1 | \mathbf{x}_i, \mathbf{w})$$

Data D: Set of sequences  $\mathbf{x}_i$  with associated interaction indicators  $R_i \in \{0, 1\}$ 

Model predicts an interaction  $R_i$  given the sequence  $\mathbf{x}_i$ :

$$y(\mathbf{x}_i, \mathbf{w}) = P(R_i = 1 | \mathbf{x}_i, \mathbf{w})$$
$$P(D|\mathbf{w}) = \prod_i y(\mathbf{x}_i, \mathbf{w})^{R_i} [1 - y(\mathbf{x}_i, \mathbf{w})]^{(1-R_i)}$$
$$\log P(D|\mathbf{w}) = \sum_i R_i \log y(\mathbf{x}_i, \mathbf{w}) + (1 - R_i) \log y(\mathbf{x}_i, \mathbf{w})$$

Maximum likelihood:  $\operatorname{argmax}_{\mathbf{w}} P(D|\mathbf{w})$ 

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Iterative optimisation scheme with gradient descent:

$$E_D(\mathbf{w}) = -\log P(D|\mathbf{w})$$
  
 $\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w})$ 

Parameter estimation Maximum likelihood:  $\operatorname{argmax}_{\mathbf{w}} P(D|\mathbf{w})$ Iterative optimisation scheme with gradient descent:  $E_D(\mathbf{w}) = -\log P(D|\mathbf{w})$  $\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w})$ Problem: overfitting!

Parameter estimation Maximum likelihood:  $\operatorname{argmax}_{\mathbf{w}} P(D|\mathbf{w})$ Iterative optimisation scheme with gradient descent:  $E_D(\mathbf{w}) = -\log P(D|\mathbf{w})$  $\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w})$ Problem: overfitting! **Regularisation**:  $P(\mathbf{w}) = \frac{1}{Z} \exp(-\alpha E_R(\mathbf{w}))$ 

Maximum a posteriori:

 $\operatorname{argmax}_{\mathbf{w}} P(\mathbf{w}|D)$ Bayes rule:  $P(\mathbf{w}|D) \propto P(D|\mathbf{w})P(\mathbf{w})$  $\operatorname{argmax}_{\mathbf{w}} [\log P(D|\mathbf{w}) + \log P(\mathbf{w})]$ 

Maximum a posteriori:  $\operatorname{argmax}_{\mathbf{w}} P(\mathbf{w}|D)$ Bayes rule:  $P(\mathbf{w}|D) \propto P(D|\mathbf{w})P(\mathbf{w})$  $\operatorname{argmax}_{\mathbf{w}}[\log P(D|\mathbf{w}) + \log P(\mathbf{w})]$  $P(\mathbf{w}) \propto \exp[-\alpha E_R(\mathbf{w})]$  $P(D|\mathbf{w}) \propto \exp[-E_D(\mathbf{w})]$  $\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w}) - \alpha \nabla_{\mathbf{w}} E_R(\mathbf{w})$ 

Weight decay:

$$\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w}) - \alpha \nabla_{\mathbf{w}} E_R(\mathbf{w})$$

Gaussian prior:  $E_R(\mathbf{w}) = \mathbf{w}^2$ 

Laplacian prior: 
$$E_R(\mathbf{w}) = |\mathbf{w}|$$

## Justification of regularisation

$$P(R = 1|S_1, S_2, \dots, S_N) = \frac{P(S_1, S_2, \dots, S_N | R = 1) P(R = 1)}{P(S_1, S_2, \dots, S_N)}$$

$$= \frac{P(S_1, S_2, \dots, S_N | R = 1) P(R = 1)}{P(S_1, S_2, \dots, S_N | R = 0) P(R = 0) + P(S_1, S_2, \dots, S_N | R = 1) P(R = 1)}$$

$$= \left(1 + \frac{P(R = 0) P(S_1, S_2, \dots, S_N | R = 0)}{P(R = 1) P(S_1, S_2, \dots, S_N | R = 1)}\right)^{-1}$$

$$= \left(1 + \left[\frac{P(R = 1)}{P(R = 0)} \frac{1}{(N - W + 1)} \sum_{m=0}^{N - W} \prod_{k=1}^{W} \frac{\psi_k(S_{m+k})}{\theta_0(S_{m+k})}\right]^{-1}\right)^{-1}$$

Define: 
$$w_k(l) = \log \frac{\psi_k(l)}{\theta_0(l)}$$
,  $w_0 = \log \frac{P(R=1)}{P(R=0)}$ 



Weight decay:  $\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w}) - \alpha \nabla_{\mathbf{w}} E_R(\mathbf{w})$ 

Gaussian prior:  $E_R(\mathbf{w}) = \mathbf{w}^2$ 

Laplacian prior: 
$$E_R(\mathbf{w}) = |\mathbf{w}|$$

The hyperparameter  $\alpha$  can be integrated out analytically

$$P(\mathbf{w}) = \int_0^\infty P(\mathbf{w}|\alpha) P(\alpha) d\alpha$$

$$P(\mathbf{w}|\alpha) = \frac{\exp(-\alpha E_R)}{Z(\alpha)}$$
$$Z(\alpha) \propto \left(\frac{1}{\alpha}\right)^W$$

where W is the dimension of  $\mathbf{w}$  (number of weights).

$$P(\mathbf{w}) = \int_0^\infty P(\mathbf{w}|\alpha) P(\alpha) d\alpha$$

Scale parameter: uninformative prior  $P(\alpha) \propto \frac{1}{\alpha}$ 

$$P(\mathbf{w}) = \int_0^\infty P(\mathbf{w}|\alpha) P(\alpha) d\alpha$$
  
=  $C \int_0^\infty \exp(-\alpha E_R) \alpha^{W-1} d\alpha$   
=  $C E_R^{-W} \int_0^\infty \exp(-\alpha E_R) (\alpha E_R)^{(W-1)} d(\alpha E_R)$   
=  $C E_R^{-W} \int_0^\infty \exp(-u) u^{W-1} du$   
=  $C E_R(\mathbf{w})^{-W} \Gamma(W)$ 

$$\log P(\mathbf{w}) = -W \log E_R(\mathbf{w}) + const$$
$$\nabla_{\mathbf{w}} \log P(\mathbf{w}) = -\frac{W}{E_R} \nabla_{\mathbf{w}} E_R(\mathbf{w})$$

Weight decay:

$$\Delta \mathbf{w} \propto -\nabla_{\mathbf{w}} E_D(\mathbf{w}) - \tilde{\alpha} \nabla_{\mathbf{w}} E_R(\mathbf{w}); \quad \tilde{\alpha} = \frac{W}{E_R}$$

Gaussian prior:  $E_R(\mathbf{w}) = \mathbf{w}^2$ Laplacian prior:  $E_R(\mathbf{w}) = |\mathbf{w}|$ 

## Peter Williams (1995)

Bayesian regularisation and pruning using a Laplacian prior

Neural Computation 7, 117–143

# Evaluation



True network

Predicted network

#### Probabilistic inference



True network

Predicted network

## Thresholding



True network

Predicted network

## Thresholding

True positives

False positives













SH3 yeast two-hybrid interaction network
Tong et al. (2002), Science 295, 321-324
285 interactions between 28 SH3 proteins and 143 binding peptides

9 binding partners per SH3 domain on average



SH3 domain protein interaction network in S. cerevisiae; from Tong et al. (2002)

No regularisation


With regularisation



## N-fold crossvalidation











Tong et al. (2002), Science 295, 321-324. SH3 domain proteins in *Saccharomyces cerevisiae*. **Yeast two-hybrid** interaction network 285 interactions between 28 SH3 proteins and 143 binding peptides Phage display interaction network 394 interactions between 28 SH3 proteins and 178 binding peptides

## Models compared in our study

- Generative model of Reiss
- Discriminative model, informative initialisation
- Ensemble of discriminative models, random initialisations



#### AUROC scores

Model	$\rightarrow$	Generative	Discriminative,	Discriminative,
		(Reiss et al.)	informative init	ensemble
Yeast	AUROC	0.61	0.67	0.67



#### AUROC scores

Model	$\rightarrow$	Generative	Discriminative,	Discriminative,
		(Reiss et al.)	informative init	ensemble
Yeast	AUROC	0.61	0.67	0.67
Phage	AUROC	0.69	0.83	0.71

# **Biological validation**

400 highest scoring interactions





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Enrichment for higher in silico scores, filter for noisy high-throughput data.

### Summary

- High-throughput interactomic data are noisy
  → Complement data with in silico predictions.
- Generative probabilistic model of Reiss & Schwikowski (2004): Several user-defined tuning parameters
- Discriminative probabilistic model of Segal et al. (2003): Overfitting
- Regularisation with Laplacian prior (Williams 1995).