HMM Algorithms

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HMM algorithms

- HMM recap
- HMM algorithms
 - Likelihood computation (forward algorithm)
 - Finding the most probable state sequence (Viterbi algorithm)
 - Estimating the parameters (forward-backward and EM algorithms)

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Warning: the maths continues!

Recap: the HMM



- A generative model for the sequence $X = (x_1, \ldots, x_T)$
- Discrete states q_t are unobserved
- q_{t+1} is conditionally independent of q_1, \ldots, q_{t-1} , given q_t
- Observations x_t are conditionally independent of each other, given q_t.

The three-state left-to-right topology for phones:



Joint likelihood of X and $Q = (q_1, \ldots, q_T)$:

$$P(X, Q|\lambda) = P(q_1)P(\mathbf{x}_1|q_1)P(q_2|q_1)P(\mathbf{x}_2|q_2)\dots$$
(1)
= $P(q_1)P(\mathbf{x}_1|q_1)\prod_{t=2}^{T}P(q_t|q_{t-1})P(\mathbf{x}_t|q_t)$ (2)

 $P(q_t)$ denotes the initial occupancy probability of each state

The parameters of the model, λ , are given by:

- Transition probabilities $a_{kj} = P(q_{t+1} = j | q_t = k)$
- Observation probabilities $b_j(\mathbf{x}) = P(\mathbf{x}|q=j)$

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Likelihood Determine the overall likelihood of an observation sequence X = (x₁,..., x_t,..., x_T) being generated by a known HMM topology, *M*.

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 → the *forward algorithm*
- Oecoding and alignment Given an observation sequence and an HMM, determine the most probable hidden state sequence → the Viterbi algorithm
- Training Given an observation sequence and an HMM, learn the state occupation probabilities, in order to find the best HMM parameters λ = {{a_{jk}}, {b_j()}}
 → the *forward-backward* and *EM* algorithms

By the HMM topology, \mathcal{M} , we can mean:

- A restricted left-to-right topology based on a known word/sentence, leading to a "trellis-like" structure over time
- A much less restricted topology based on a grammar or language model – or something in between
- The forward/backward algorithms are not (generally) suitable for unrestricted topologies

Example: trellis for /k ae t/



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• How many paths Q do we have to calculate?

$$\sim \underbrace{N \times N \times \cdots N}_{T \text{ times}} = N^T \qquad N: \text{ number of HMM states} \\ T: \text{ length of observation}$$

e.g. $N^{T} \approx 10^{10}$ for N = 3, T = 20

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• Computation complexity of multiplication: $O(2TN^{T})$

Likelihood: The Forward algorithm

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- Rather than enumerating each sequence, compute the probabilities recursively (exploiting the Markov assumption)
- Reduces the computational complexity to $O(TN^2)$
- Visualise the problem as a *state-time trellis*



The forward probability

Define the *Forward probability*, $\alpha_t(j)$: the probability of observing the observation sequence $\mathbf{x}_1 \dots \mathbf{x}_t$ and being in state j at time t:

$$\alpha_j(t) = p(\mathbf{x}_1, \ldots, \mathbf{x}_t, q_t = j | \mathcal{M})$$

We can recursively compute this probability

Initial and final state probabilities

It what follows it is convenient to define:

• an additional single initial state $S_I = 0$, with transition probabilities

$$a_{0j} = P(q_1 = j)$$

denoting the probability of starting in state j

- a single final state, S_E, with transition probabilities a_{jE} denoting the probability of the model terminating in state j.
- S_I and S_E are both non-emitting

1. Likelihood: The Forward recursion

Initialisation

$$\begin{array}{ll} \alpha_j(0) = 1 & \quad j = 0 \\ \alpha_j(0) = 0 & \quad j \neq 0 \end{array}$$

Recursion

$$\alpha_j(t) = \sum_{i=1}^J \alpha_i(t-1)a_{ij}b_j(\mathbf{x}_t) \qquad 1 \le j \le J, \ 1 \le t \le T$$

Termination

$$p(\mathbf{X} | \mathcal{M}) = \alpha_E = \sum_{i=1}^J \alpha_i(T) a_{iE}$$

 s_I : initial state, s_E : final state

1. Likelihood: Forward Recursion

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- If we are performing decoding or forced alignment, then only the most likely path is needed
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- We need to keep track of the states that make up this path by keeping a sequence of *backpointers* to enable a Viterbi *backtrace*: the backpointer for each state at each time indicates the previous state on the most probable path





Backpointers to the previous state on the most probable path



2. Decoding: The Viterbi algorithm

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• Termination

$$P^* = V_E = \max_{i=1}^{J} V_T(i) a_{iE}$$
$$s_T^* = B_E = \arg \max_{i=1}^{J} V_i(T) a_{iE}$$

Backtrace to find the state sequence of the most probable path



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- Parameters \mathcal{M} :
 - Transition probabilities *a_{ij}*:

$$\sum_{j}a_{ij}=1$$

 Gaussian parameters for state *j*: mean vector μ_j; covariance matrix Σ_j

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• Likewise if Z_j is the set of observed acoustic feature vectors assigned to state j, we can use the standard maximum likelihood estimates for the mean and the covariance:

$$\hat{\mu}_j = rac{\sum_{oldsymbol{x}\in Z_j}oldsymbol{x}}{|Z_j|} \hat{\Sigma}_j = rac{\sum_{oldsymbol{x}\in Z_j}(oldsymbol{x}-\hat{\mu}_j)(oldsymbol{x}-\hat{\mu}_j)^T}{|Z_j|}$$

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Compare with component occupation probability in a GMM

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E-step estimate the state occupation probabilities (Expectation)

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 - E-step estimate the state occupation probabilities (Expectation)
 - M-step re-estimate the HMM parameters based on the estimated state occupation probabilities (Maximisation)

• To estimate the state occupation probabilities it is useful to define (recursively) another set of probabilities—the *Backward probabilities*

$$\beta_j(t) = p(\mathbf{x}_{t+1}, \ldots, \mathbf{x}_T | q_t = j, \mathcal{M})$$

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Recursion

$$eta_i(t) = \sum_{j=1}^J a_{ij} b_j(\mathbf{x}_{t+1}) eta_j(t+1) \quad ext{for } t = T-1, \dots, 1$$

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Termination

$$p(\mathbf{X} \mid \mathcal{M}) = \beta_0(0) = \sum_{j=1}^J a_{0j} b_j(\mathbf{x}_1) \beta_j(1) = \alpha_E$$

Backward Recursion

$$\beta_j(t) = p(\mathbf{x}_{t+1}, \dots, \mathbf{x}_T \mid q_t = j, \mathcal{M}) = \sum_{j=1}^J a_{ij} b_j(\mathbf{x}_{t+1}) \beta_j(t+1)$$



State Occupation Probability

- The state occupation probability γ_j(t) is the probability of occupying state j at time t given the sequence of observations
- Express in terms of the forward and backward probabilities:

$$\gamma_j(t) = P(q_t = j | \mathbf{X}, \mathcal{M}) = \frac{1}{\alpha_E} \alpha_j(t) \beta_j(t)$$

recalling that $p(\mathbf{X} | \mathcal{M}) = \alpha_E$

Since

$$\begin{aligned} \alpha_j(t)\beta_j(t) &= p(\mathbf{x}_1, \dots, \mathbf{x}_t, q_t = j \mid \mathcal{M}) \\ p(\mathbf{x}_{t+1}, \dots, \mathbf{x}_T \mid q_t = j, \mathcal{M}) \\ &= p(\mathbf{x}_1, \dots, \mathbf{x}_t, \mathbf{x}_{t+1}, \dots, \mathbf{x}_T, q_t = j \mid \mathcal{M}) \\ &= p(\mathbf{X}, q_t = j \mid \mathcal{M}) \end{aligned}$$

$$P(q_t = j | \mathbf{X}, \mathcal{M}) = \frac{p(\mathbf{X}, q_t = j | \mathcal{M})}{p(\mathbf{X} | \mathcal{M})}$$

- The sum of state occupation probabilities through time for a state, may be regarded as a "soft" count
- We can use this "soft" alignment to re-estimate the HMM parameters:

$$\hat{\mu}_{j} = \frac{\sum_{t=1}^{T} \gamma_{j}(t) \mathbf{x}_{t}}{\sum_{t=1}^{T} \gamma_{j}(t)}$$
$$\hat{\Sigma}_{j} = \frac{\sum_{t=1}^{T} \gamma_{j}(t) (\mathbf{x}_{t} - \hat{\mu}_{j}) (\mathbf{x} - \hat{\mu}_{j})^{T}}{\sum_{t=1}^{T} \gamma_{j}(t)}$$

Re-estimation of transition probabilities

• Similarly to the state occupation probability, we can estimate $\xi_{i,j}(t)$, the probability of being in *i* at time *t* and *j* at t + 1, given the observations:

$$\xi_t(i, j) = P(q_t = i, q_{t+1} = j | \mathbf{X}, \mathcal{M})$$
$$= \frac{p(q_t = i, q_{t+1} = j, \mathbf{X} | \mathcal{M})}{p(\mathbf{X} | \mathcal{M})}$$
$$= \frac{\alpha_i(t) a_{ij} b_j(\mathbf{x}_{t+1}) \beta_j(t+1)}{\alpha_E}$$

• We can use this to re-estimate the transition probabilities

$$\hat{a}_{ij} = rac{\sum_{t=1}^{T} \xi_{i,j}(t)}{\sum_{k=1}^{J} \sum_{t=1}^{T} \xi_{i,k}(t)}$$

Pulling it all together

• Iterative estimation of HMM parameters using the EM algorithm. At each iteration

E step For all time-state pairs

- Recursively compute the forward probabilities $\alpha_i(t)$ and backward probabilities $\beta_i(t)$
- 2 Compute the state occupation probabilities $\gamma_j(t)$ and $\xi_{i,j}(t)$
- M step Based on the estimated state occupation probabilities re-estimate the HMM parameters: mean vectors μ_j , covariance matrices Σ_j and transition probabilities a_{ij}
- The application of the EM algorithm to HMM training is sometimes called the Forward-Backward algorithm or Baum-Welch algorithm

Extension to a corpus of utterances

- We usually train from a large corpus of R utterances
- If \mathbf{x}_t^r is the *t* th frame of the *r* th utterance \mathbf{X}^r then we can compute the probabilities $\alpha_j^r(t)$, $\beta_j^r(t)$, $\gamma_j^r(t)$ and $\xi_{i,j}^r(t)$ as before
- The re-estimates are as before, except we must sum over the *R* utterances, eg:

$$\hat{\boldsymbol{\mu}}_{j} = \frac{\sum_{r=1}^{R} \sum_{t=1}^{T} \gamma_{j}^{r}(t) \boldsymbol{x}_{t}^{r}}{\sum_{r=1}^{R} \sum_{t=1}^{T} \gamma_{j}^{r}(t)}$$

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 In addition, we usually employ "embedded training", in which fine tuning of phone labelling with "forced Viterbi alignment" or forced alignment is involved. (For details see Section 9.7 in Jurafsky and Martin's SLP)

- The assumption of a Gaussian distribution at each state is very strong; in practice the acoustic feature vectors associated with a state may be strongly non-Gaussian
- In this case an *M*-component Gaussian mixture model is an appropriate density function:

$$b_j(\mathbf{x}) = p(\mathbf{x} | q = j) = \sum_{m=1}^M c_{jm} \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}_{jm}, \boldsymbol{\Sigma}_{jm})$$

Given enough components, this family of functions can model any distribution.

• Train using the EM algorithm, in which the component estimation probabilities are estimated in the E-step

EM training of HMM/GMM

 Rather than estimating the state-time alignment, we estimate the component/state-time alignment, and component-state occupation probabilities γ_{jm}(t): the probability of occupying mixture component m of state j at time t.

 $(\xi_{tm}(j) \text{ in Jurafsky and Martin's SLP})$

• We can thus re-estimate the mean of mixture component *m* of state *j* as follows

$$\hat{\mu}_{jm} = rac{\sum_{t=1}^{T} \gamma_{jm}(t) \mathbf{x}_t}{\sum_{t=1}^{T} \gamma_{jm}(t)}$$

And likewise for the covariance matrices (mixture models often use diagonal covariance matrices)

• The mixture coefficients are re-estimated in a similar way to transition probabilities:

$$\hat{c}_{jm} = \frac{\sum_{t=1}^{I} \gamma_{jm}(t)}{\sum_{m'=1}^{M} \sum_{t=1}^{T} \gamma_{jm'}(t)}$$

- The forward, backward and Viterbi recursions result in a long sequence of probabilities being multiplied
- This can cause floating point underflow problems
- In practice computations are performed in the log domain (in which multiplies become adds)
- Working in the log domain also avoids needing to perform the exponentiation when computing Gaussians

- HMMs provide a generative model for statistical speech recognition
- Three key problems
 - Computing the overall likelihood: the Forward algorithm
 - 2 Decoding the most likely state sequence: the Viterbi algorithm
 - Estimating the most likely parameters: the EM (Forward-Backward) algorithm
- Solutions to these problems are tractable due to the two key HMM assumptions
 - Conditional independence of observations given the current state
 - Markov assumption on the states

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